

Islamistic terror,	, the war on	Iraq and t	the job	prospects	of Arab	men in	Britain
Braakmann, Nils							

Publication date: 2007

Document Version Publisher's PDF, also known as Version of record

Link to publication

Citation for pulished version (APA):

Braakmann, N. (2007). *Islamistic terror, the war on Iraq and the job prospects of Arab men in Britain: Does a country's direct involvement matter?* (Working paper series in economics; No. 70). Institut für Volkswirtschaftslehre der Universität Lüneburg.

Copyright and moral rights for the publications made accessible in the public portal are retained by the authors and/or other copyright owners and it is a condition of accessing publications that users recognise and abide by the legal requirements associated with these rights.

- Users may download and print one copy of any publication from the public portal for the purpose of private study or research.
  You may not further distribute the material or use it for any profit-making activity or commercial gain
  You may freely distribute the URL identifying the publication in the public portal?

Take down policy
If you believe that this document breaches copyright please contact us providing details, and we will remove access to the work immediately and investigate your claim.

Download date: 27. Apr.. 2024

# Islamistic Terror, the War on Iraq and the Job Prospects of Arab Men in Britain: Does a Country's Direct Involvement matter?

**NORKING** 

by Nils Braakmann

University of Lüneburg Working Paper Series in Economics

No. 70

December 2007

www.leuphana.de/vwl/papers ISSN 1860 - 5508

# Islamistic Terror, the War on Iraq and the Job Prospects of Arab Men in Britain: Does a Country's Direct Involvement Matter?

Nils Braakmann\* Leuphana University Lüneburg

This version: December 20, 2007

#### Abstract

This paper examines whether the labor market prospects of Arab men in England are influenced by recent Islamistic terrorist attacks and the war on Iraq. We use data from the British Labour Force Survey from Spring 2001 to Winter 2006 and treat the terrorist attacks on the USA on September 11th, 2001, the Madrid train bombings on March 11th, 2004 and the London bombings on July 7th, 2005, as well as the beginning of the war on Iraq on March 20th, 2003, as natural experiments possibly having led to a change in attitudes toward Arab or Muslim men. Using treatment group definitions based on ethnicity, country of birth, current nationality, and religion, evidence from regression-adjusted difference-in-differences-estimators indicates that the real wages, hours worked and employment probabilities of Arab men were unchanged by the attacks. This finding is in line with prior evidence from Europe.

**Keywords:** Discrimination, September 11th, Islamistic terror, employment, wages **JEL Classification:** J71, J79

<sup>\*</sup>Empirical Economics, Institute of Economics, Leuphana University Lueneburg, braak-mann@uni.leuphana.de, Tel.: 0049 (0) 4131 677 2303, Fax: 0049 (0) 4131 677 2026 The author would like to thank Joachim Wagner for helpful hints and overall support. All calculations were performed using Stata 9.2 SE (StataCorp 2005). All do-files are available from the author on request.

## 1 Introduction

Following the terrorist attacks on September 11th, 2001 a number of studies have been concerned with the (economic) causes (e.g. Krueger and Malečková 2003, Abadie 2006, Piazza 2006, Krueger and Laitin 2007) or consequences (e.g. Abadie and Gardeazabal 2003, 2007, Abadie and Dermisi 2006, Frey, Luechinger and Stutzer 2007) of terrorism. In that literature a small but growing number of papers have been concerned with the economic consequences of the 9/11-attacks for Arabs or Muslims living in western countries.<sup>1</sup>

Directly after the attacks a number of reports collected by various organizations suggested a rise in discrimination and hostility toward persons perceived to be Arabs or Muslim (see Allen and Nielsen (2002) for Europe and the Arab American Institute (2003) and the Arab-American Anti-Discrimination Committee (2003) for the US). Up to today, four studies have investigated whether this anecdotal evidence was accompanied by observable changes in the labor market prospects of Arabs or Muslims. A short overview of these studies can be found in table 1, a detailed description is provided in the following paragraphs.

(Table 1 about here.)

For the US, Dávila and Mora (2005) use data from the American Community Surveys and focus on the wages of younger men between 25 and 40 years of age. Using linear and quantile regression as well as decomposition techniques, they find that the wages of men from Middle Eastern countries have been harmed most by the attacks, while less of an impact could be found for African Arabs and other Arabs relative to US-born non-Hispanics.

Also focusing on the US, Kaushal, Kaestner, and Reimers (2007) use regression-adjusted difference-in-differences-estimates on Current Population Survey data to asses changes in job prospects and mobility for persons from predominantly Muslim / Arab countries relative to natives and other migrants. Their results indicate that the real wages and weekly earnings of Arab men were reduced by an amount of 9-11% as a consequence of the attacks, though this effect seems to have been temporarily with a significant rebound noted in 2005. Furthermore, they find hints that intrastate mobility of Arab men was also reduced by the 9/11-attacks, while employment and hours worked seem to have been relatively uninfluenced by the attacks.

For Europe, Aslund and Rooth (2005) focus on exits from unemployment for men in Sweden. They use difference-in-differences-estimators on administrative labor market data and look at the development of employment prospects of those from the Middle East relative to a number of control groups. Their findings indicate that

<sup>&</sup>lt;sup>1</sup>There has also been some interest in the question whether other immigrant groups have been harmed in the aftermath of the 9/11-attacks, see Orrenius and Zavodny (2006).

there has been no significant drop in re-employment probabilities for persons from the Middle East compared to natives, people from the Nordic countries and from former Yugoslavia, Western and Eastern Europeans, Latin Americans, Asians, and Africans that could be attributed to the 9/11-attacks.

In a similar study for Germany, Braakmann (2007) also uses regression-adjusted difference-in-differences-estimators on administrative data from the Federal Employment Agency and Social Security. He uses various treatment and control group definitions as well as a number of robustness checks. His findings confirm the results from the study by Åslund and Rooth (2005), namely that the employment prospects of Arabs do no seem to have been harmed by the attacks.

Unfortunately, the picture that emerges from these studies is far from clear. At a first glance, there seems to be a difference between the US and Europe. However, the factors driving the different results from these studies are not entirely obvious. A first possible explanation is that the differences between Europe and the US might reflect genuine differences in the respective population's change in attitudes towards Arabs or Muslims. The apparently stronger reaction of the US population to the attacks could then be related to the fact that the US were the direct target of the 9/11-attacks. While this explanation seems intuitively appealing, there might be other explanations: Firstly, there might be intervening factors like differences in labor market institutions. Both Sweden and Germany are highly regulated and institutionalized labor markets which would be expected to weaken the impact of a possible change in attitudes. Secondly, one might raise the question whether a difference between the US and Europe in fact exists: Both European studies focus only on employment probabilities – a variable where Kaushal, Kaestner and Reimers (2007) could also not find an impact for the US.

This study attempts to address these issues by using British labor market data for the years 2000 to 2006. First, we will look at the same labor market outcomes as Kaushal, Kaestner and Reimers (2007) – real weekly and hourly wages, hours worked and the probability of being employed – thus providing better comparative evidence between Europe and the US. Secondly, to test the idea that a country's direct involvement in terrorist attacks might matter for explaining changes in the job prospects of Arabs or Muslims, we will use the fact that Britain was hit directly by a terrorist attack on July 7th, 2005. Finally, looking at evidence from one country over time allows us to rule out the possibility that differences in labor market institutions interfere with the results.

More specifically, we will use the terrorist attacks on the US on September 11th, 2001, the Madrid train bombings on March 11th, 2004 and the London bombings on July 7th, 2005, as well as the beginning of the war on Iraq on March 20th, 2003, as natural experiments and use regression-adjusted difference-in-differences-estimators to assess the change in the four labor market outcomes outlined above. The data used comes from the British Quarterly Labour Force Survey from Spring 2001 to Winter 2006 and allows us to gain further insight into the (possible) discrimination mechanism at work as it provides the possibility to use different definitions for

the treatment and control groups based on self-assessed ethnicity, country of birth, current nationality and (beginning in 2002) religion.

The rest of this paper is organized as follows: Section 2 provides some background information on the events from September 2001 to July 2005. The data used is described in section 3 along with some descriptive information on the sample used. Section 4 contains a description on the econometric model whose results are found in section 5. Finally, section 6 concludes.

# 2 Background

This section presents a short chronological review of the events relevant in the context of this paper. Note that this paper does not attempt to provide an in-depth description of the time following the 9/11-attacks and in particular does not attempt to deal with the highly controversial questions surrounding the war on Iraq and the Madrid train bombings. Since this paper focuses on the impact of these events in England, we will try to provide specific information for this country where possible.

The timeline starts at September 11th, 2001 when three airplanes were flown into the World Trade Center and the Pentagon with a fourth crashing on a field in Pennsylvania.<sup>2</sup> In the months directly following the attacks British media reported a rise in violent acts against Muslims and Mosques. British politicians and other public persons, including Prime Minister Tony Blair, the Prince of Wales and the Archbishop of Canterbury, issued calls for calm, tolerance, and differentiation. An exemption from these calming voices was the right-wing British National Party that launched an islamophobic campaign after the 9/11-attacks that displayed Christianity as being threatened by Islam. Media response seems to have been mixed with rather much time being devoted to more radical Muslims (see Allen and Nielsen 2002, pp. 29-30 for further details). In a retrospective study conducted among British Muslims between October and December 2002 that focused on differences before and after 9/11, Sheridan (2006) finds that 82.6% of the respondents reported an increase in implicit racism and religious discrimination, while 76.3% reported an increase in "general discriminatory experiences".

The 20th of March 2003 marked the beginning of the (highly disputed) war on Iraq, led by American and British Forces, that was officially aimed at finding and destroying Iraqi weapons of mass destruction, as well as severing suspected links between the Iraqi government and islamistic terrorists. In May 2003 US President George Bush announced the official end of the war, however, up to today a large number of American and British soldiers remain in Iraq securing a more or less fragile peace.

On March 11th, 2004 the first mayor terrorist attack by islamistic terrorists after 9/11 occurred when several bombs exploded in commuter trains in Madrid

 $<sup>^2</sup>$ An in-depth description of the events on September 11th, 2001 can be found in the official report by the "National Commission on Terrorist Attacks Upon the United States" (2004).

during the morning rush hour. The attack that was initially blamed on the Basque separatist group ETA by the then ruling government under Prime Minister Aznar caused the deaths of 191 people and the injuries of over 2,000. The events on this day and their handling by the conservative government is believed to have had a mayor impact on the outcome of the Spanish election on March 14th, 2004 that brought the Socialist Workers Party under José Zapatero to power (for a recent analysis of the attack's impact on the electoral outcomes see Montalvo 2006).

On July 7th, 2005 England was hit directly by a terrorist attack, when a group of four young British-born Muslims set off four bombs in the London underground and one double-decker bus resulting in the deaths of 56 people (including the four attackers) and several hundred injuries. A similar attack on July 21st, 2005 resulted in no casualties as only the detonators of the bombs exploded.

Shortly after the attacks government officials and the police issued warnings that violent reactions against the Muslim community would not be tolerated and tried to ensure that a distinction was made between the terrorists and Muslims as a whole. Nevertheless, several organizations reported a rise in incidents against Muslims and Muslim organizations issued warnings of possible negative reactions and threats, while at the same time condemning the attacks. Media responses were generally balanced with a shift towards integration issues and a possible radicalization in the Muslim communities after it became known that the attacks were conducted by British-born Muslims (for a detailed account of the events, including a full timeline and a more detailed description of the reactions in the United Kingdom, see European Monitoring Centre on Racism and Xenophobia 2005).

### 3 Data

The data used in this study comes from the British (Quarterly) Labour Force Survey (LFS), a survey conducted among households living at private addresses or National Health Service accommodations in the UK by the Office of National Statistics since 1973.<sup>3</sup> The survey is collected quarterly since Spring 1992. From 1992 to May 2006 data collection took place in a seasonal pattern with surveys being conducted in winter (December to February), spring (March to May), summer (June to August) and autumn (September to November). Due to EU regulations the LFS moved to calendar quarters in May 2006 with surveys now covering the periods January to March, April to June, July to September and October to December.

The current sample size is approximately 50,000 responding households in Great Britain with an additional 2,000 being added from Northern Ireland resulting in a coverage of 0.1% of the target population. Each household is surveyed in five consec-

<sup>&</sup>lt;sup>3</sup>The data can be accessed free of charge at the "Economic and Social Data Service" after registration. See http://www.esds.ac.uk/ for further information. To facilitate replication all Stata do-files used in the analysis can be obtained from the author.

utive quarters in a rotating panel design. Since roughly one fifth of the respondents enter and leave each quarter there is an 80% overlap between two adjacent quarters.

The survey is designed to provide information on the labor market status and personal situation of individuals living in the UK during a reference period, usually a specific week. The questionnaire therefore encompasses information on employment, including information on the current employer, socio-demographic characteristics, education, and wages as well as information on the respective household. Most importantly for the scope of this paper, the data contains information on a respondent's ethnicity, country of birth, current nationality and religion which can be used to construct treatment and control groups. Additionally, the data provides information on a number of relevant labor market outcomes and control variables. As the data contains information on the timing of the interview and the reference week the information relates to it is also possible to assess whether a specific individual was observed before or after any of the events of interest.

In this paper four different definitions of "Arabs" or "Muslims" are considered. The first definition is based on self-assessed ethnicity, where those individuals reporting a "Pakistani" or "Bangladeshi" ethnicity form the treatment group. This definition is in line with findings on Islamophobia by the European Monitoring Centre on Racism and Xenophobia (2006, p.17) stating that Pakistanis and Bangladeshis have the highest risk of being victim of a racially motivated crime. Additionally, the majority of (migrant) Muslims in the UK originates from those two countries (European Monitoring Centre on Racism and Xenophobia 2006, p. 22). As comparison groups we will look at individuals with a British ethnicity and those reporting any other non-white ethnicity.

The second definition is based on an individual's country of birth. The treatment group is formed by those individuals born in Algeria, Bangladesh, Egypt, Pakistan, and all Middle East countries with the exception of Israel. As controls we use individuals born in Britain and those born in southern Africa, Asia, South America, and the Caribbean. The third definition uses the same countries, but is based on current nationality rather than country of birth.

Note that we follow Kaushal, Kaestner, and Reimers (2007) in excluding Indians from the comparison groups formed by these definitions. First, Indians might be expected to look somewhat similar to Pakistani and Bangladeshi as these stem from almost the same region. Secondly, while the Indian population is predominantly Hindu, there is a strong Muslim minority which might be expected to underlie the same discriminational mechanism as the treatment group.

Finally, for the fourth definition we rely on religion. The treatment group is formed by Muslims who are being compared with Christians and other religions respectively. Note that Muslims can only be identified in the data from Spring 2002 onward which means that we can only measure the impact of the beginning of the war on Iraq and the Madrid and London bombings relative to the time after 9/11. Note further that Sikhs are excluded from the control group "other religions" as

several reports (e.g. Allen and Nielsen 2002) suggest that these were often confused with Muslims.

The data used in this paper covers the time from Spring 2001, when a major revision of the ethnicity question was introduced, to Winter 2006. To minimize the impact of regional differences we focus on persons living in England and consequently exclude households with residence in Scotland, Wales and Northern Ireland. We further exclude individuals who are under 16 or over 65 years of age to restrict the sample to the working population. Finally, the estimation sample is restricted to males since case numbers for women from the treatment groups are too low to provide reliable estimates of the parameters of interest.

Tables 2 and 3 display information on some key variables of the respective samples. Note that there are both differences between the different definitions of the treatment group and between the respective treatment and control groups. A common finding over all group definitions is that the members of the respective treatment groups have somewhat less favorable labor market relevant characteristics than members of the control groups.

(Tables 2 and 3 about here.)

#### 4 Econometric Model

Before we outline the estimation strategy, several basic facts about the problem at hand should be noted. Interest in this paper lies in the estimation of the causal (treatment) effect of recent terrorist attacks and the related war on Iraq on several labor market outcomes caused by a possible shift in attitudes towards Arabs or Muslims. Note that there is a clear theoretical one-way causality between these interventions and the outcomes of interest.

Furthermore all three terrorist attacks are unexpected events that can be considered natural experiments leading to an exogenous shift in attitudes towards Arabs or Muslims. The case of the Iraq war is somewhat different: As the begin of a war is usually not completely unanticipated it might be the case that a possible change in attitudes took place at some different point in time, e.g. after going to war was first discussed in the public or after the first British casualties were reported. For the scope of this paper this fact implies that care should be taken when attaching a causal interpretation to the effect associated with the war.

Finally, selection out of or into the treatment group can be ruled out for both ethnicity and country of birth as both are strictly exogenous variables. Religion and current nationality however can be influenced by personal decisions: Persons may decide to become naturalized or convert to some other religion to avoid discrimination. However, since this paper uses all four variables to define treatment and control groups, it is possible to make a statement whether the choice of this variable influences the results.

To model the effects of the different terrorist attacks and the beginning of the war on Iraq consider a regression-adjusted differences-in-difference estimator of the form

$$y_i = \alpha + \beta' X_i + \chi * d_i + \sum_{j=1}^4 \delta_j * t_{ji} + \sum_{j=1}^4 \tau_j * (t_{ji} * d_i) + \epsilon_i$$
 (1)

where  $y_i$  is the respective outcome of interest,  $\epsilon_i$  is an error term,  $X_i$  is a matrix of control variables,  $d_i$  is a dummy variable indicating whether an individual belongs to the respective treatment group and  $t_{ji}$  indicates the period in which an individual was observed. More specifically,  $t_{ji}$  takes the values displayed in figure 1 that also gives an overview of the complete setup of the estimation.

The parameters of interest measuring the change in labor market outcomes for the treatment group after the respective event are given by  $\tau_1$  to  $\tau_4$  for the events from 9/11 to the London bombings respectively. Note that the setup of the period dummies implies that each  $\tau$  measures the impact of the associated event relative to the preceding period – that is e.g.  $\tau_3$ , the parameter associated with the Madrid bombings, measures the impact of that event relative to the situation after the begin of the war on Iraq.<sup>4</sup>

For (log) hourly wages, (log) weekly wages and weekly hours worked as dependent variables equation (1) is estimated by OLS while the probability of being employed is estimated by standard Probit regression. Note that two further sample restrictions are imposed for the OLS estimations. First, the estimations are only conducted for those in employment as wages and to some extent hours worked are ill-defined for the unemployed or those out of the labor force. Secondly, due to low case numbers in the treatment groups individuals working in agriculture, fishing, mining, private households and extraterritorial organizations are excluded from these estimations.

 $X_i$  contains information on education measured by 6 dummies, age in years and a dummy variable indicating whether the individual has health problems hindering at work, as well as occupation fixed effects based on sub mayor groups, regional fixed effects based on government regions and monthly and yearly fixed effects. In the wage and hours regressions we furthermore include tenure (measured by several dummies), firmsize (2 dummy variables) and industry group dummies.

A central assumption of the difference-in-differences approach is that both treatment and control groups would have experienced the same trend in the absence of the treatment. To assess the validity of this assumption, pseudo-interventions that is artificial events defined to have taken place one month before the actual event of interest, e.g. August 11th, 2001 for the September 11th attacks, are used. The

<sup>&</sup>lt;sup>4</sup>All calculations were also performed using a simple difference-in-differences-estimator without adjusting for possible differences in control variables. The results were not substantially different. Detailed results can be found in the appendix.

difference-in-differences-estimator is then calculated using these artificial events as the actual intervention. As the interaction terms in equation (1) measure the divergence of trends in the treatment and control groups after the respective interventions, we would expect the coefficients of the pseudo-interventions to be insignificant if the common trend assumption is valid. Note that a violation of this assumption does not necessarily invalidate the difference-in-differences-analysis: While diverging trends prior to the event of interest introduce bias in the coefficients of interest, the direction of this bias can be seen from the estimated coefficients of the pseudo-interventions. Depending on the results of the actual difference-in-differences analysis and the direction of the bias it might be possible to interpret the coefficients of the actual interventions as lower or upper bounds for the effect of interest.

(Table 4 about here.)

Table 4 gives an overview of the relevant parameters for the pseudo-interventions. Note that the common trend assumption cannot be rejected for most of the treatment/control group pairs. The exceptions will be discussed below jointly with the main results for difference-in-differences-analysis.

#### 5 Results

Consider first the information on weekly wages shown in table 5. Note that the time dummies do not seem to indicate a large impact of any of the events of interests on the population as a whole. The associated coefficients are generally either insignificant or (small) positive, the only exception being the time after 9/11 in one of the ethnicity specifications. Next, note that the dummies for the treatment group are always associated with a negative point estimate that is also significant when using country of birth (regardless of the comparison group) or comparing Muslims with Christians.

(Table 5 about here.)

Now, turn to the parameters of interest associated with the period-group-interaction terms. Remember that Kaushal, Kraemer and Reimers (2007) found – depending on the specification used – a 10 to 14 percent decline in weekly wages for Arab men after 9/11. A similar result, however, cannot be found for England: All coefficients for the interaction term associated with 9/11 are insignificant. Additionally, all coefficients have positive signs which is rather unexpected.

Basically the same results can be seen for the impact of the Iraq war on the treatment group. Here again, insignificant results are obtained for all definitions of the treatment group. The associated point estimates are always positive with the

exception of the groups defined by religion where a negative, but still insignificant, impact is found.

For the Madrid train bombings we find mostly negative point estimates that are always insignificant. Note however, that the insignificance is to some extent driven by large standard errors that may disguise an otherwise large effect.

Finally, consider the results for the London bombings. Here, all point estimates are positive and rather large in magnitude. However, all but one are also insignificant on all conventional levels. The only significant coefficient indicates that the wages of Muslims have risen by approximately 12 % relative to those of Christians after the London bombings. One should keep in mind though that this counterintuitive result may be a purely statistical effect related to the rather large number of significance tests.

Consider next the estimations for hourly wages shown in table 6. For this outcome Kaushal, Kraemer and Reimers (2007) found a 9 to 11 percent decline for Arab men in the US. Again, these results are not confirmed for the British labor market. In fact, the pattern of results is practically identical to those obtained for weekly wages thus resulting in the same conclusions as outlined above. The only substantial difference is a now significant and rather large wage penalty for those with Pakistani / Bangladeshi ethnicity relative to those with British ethnicity that could not be found for weekly wages.

#### (Table 6 about here.)

Now turn to the results for hours worked displayed in table 7. Consider first the coefficients for the time dummies. The point estimates show a general negative trend after 9/11 and no clear results for the remaining events. Additionally, the estimates are generally rather small for most of the treatment/control group combinations and insignificant in all specifications.

There is also no clear result for the coefficient associated with the treatment group dummy: Strictly negative point estimates for all comparison groups are obtained for the specification using country of birth, while the opposite result can be found when using religion. For the remaining two specifications, the results generally vary with the control groups used. However, none of these results seem to hold outside the sample with all results being insignificant on conventional levels.

#### (Table 7 about here.)

The same results hold for the coefficients of the interaction terms measuring the impact of the respective events of interest: While mostly positive point estimates are obtained for 9/11 and the London bombings, the signs of the remaining coefficients tend to vary unsystematically with the treatment/control group combination. However, none of them is significant on any conventional level. One should note

at this point that this result is similar to those obtained by Kaushal, Kraemer and Reimers (2007) who also did not find a significant impact of the 9/11 attacks on workings hours.

Finally, consider the results from the Probit estimation of the employment probability shown in table 8. Here, we find negative point estimates associated with the 9/11 attacks in all specifications. This effect is also significant on conventional levels when using the respective native groups as controls.

The results for the next two period dummies are less clear cut: The point estimates have varying signs for the periods after the beginning of the Iraq war and the Madrid train attacks. For the period following the London bombings we obtain negative point estimates over all specifications. However, none of the above mentioned effects is significant on any conventional level. The treatment group dummy in this set of estimations is always associated with a negative impact, thus indicating a lower employment probability of Arabs or Muslims compared with any other group.

#### (Table 8 about here.)

The results for the interaction terms are again mixed: For the 9/11-attacks and the begin of the Iraq war we usually obtain positive point estimates that are insignificant on all conventional levels. The only exception is found for those born in an Arab country compared with natives where the point estimate is negative though still insignificant. For the Madrid train bombings we find positive point estimates in the specifications using ethnicity and country of birth and negative point estimates when using current nationality and religion. Again non of these effects is significant on any conventional levels. Furthermore, for two of the negative point estimates, Arabs contrasted with those with a British/UK nationality in the case of current nationality and Muslims vs. other religions in the case of religion, the pseudo-interventions indicated pre-existing negative trends that may have influenced the negative point estimates. Finally, for the London bombings we obtain almost exclusively positive, though insignificant point estimates. The only exception, a highly significant positive effect found in the specification using those with an Arab nationality compared with those with a British/UK nationality, is most likely caused by a pre-existing positive trend that was found using the pseudo-interventions.

Taken together, our results imply that the job prospects of Arab men in England have not been significantly harmed by either the three mayor terrorist attacks conducted by islamistic terrorists after 2000 or by the beginning of the Iraq war. Furthermore, the results indicate that a country's direct involvement in acts of terrorism, in our case the London bombings of 2005, does not seem to cause a rise in discrimination. The latter finding is in line with a report by the European Monitoring Centre on Racism and Xenophobia (2005) that also did not find a lasting effect of the attacks.

#### 6 Conclusion

This paper uses data from the British Labour Force Survey for the years 2001 to 2006 and regression-adjusted difference-in-differences-estimators to gain further insight into the question whether islamistic terrorism is harmful for the job prospects of Arabs or Muslims living in Western countries. More specifically, this paper uses the fact that England was hit "indirectly" by the attacks on the Pentagon and the World Trade Center in 2001 and the Madrid trains bombings in 2004 and "directly" by the bomb attacks in London in 2005 to provide an answer to the question whether a country's direct involvement in acts of terrorism influences the labor market prospects of those possibly associated with the terrorists. Furthermore, this paper uses for the first time the same outcome variables as previous US studies thus allowing to decide whether differences between the US and Europe found in these studies can be explained by different choices of the dependent variables. Finally, this paper is the first to use treatment group definitions based on more than one variable, in this case ethnicity, country of birth, current nationality and religion, to gain further insight into the question against which group a possible discrimination is directed.

Our results indicate that neither of the attacks influenced the wages, the working hours or the employment probability of Arab or Muslim men in England. In particular, the fact that the labor market prospects of Arabs remain unchanged after the London bomb attacks indicate that a country's direct involvement in acts of terrorism does – at least in this particular case – not seem to have a large impact on the discrimination of Arabs. This result is stable over all definitions of the treatment and control groups used.

It also confirms the evidence from prior studies for Sweden (Åslund and Rooth 2005) and Germany (Braakmann 2007) that found no evidence for an increase in discrimination after the terrorist attacks on September 11th, 2001. Furthermore, it is in line with the reports from the European Monitoring Center on Racism and Xenophobia (2005, 2006) that pointed towards no (lasting) impacts of the terrorist attacks.

Regarding the differences between the US and Europe found in previous studies, the results indicate that these differences were not merely a result of the choice of different outcomes. Furthermore, the differences cannot solely be explained by the fact that the US were the (only) direct target of the attacks. However, a possible explanation for the apparently much stronger reaction in the US might be the different scale of the attacks. While the attacks in London (and Madrid) were some of the largest terrorist attacks in Europe, both were smaller than 9/11 in terms of casualties and none of them had the massive impact on the public opinion that 9/11 had.

# 7 Bibliography

- 1. Abadie, Alberto, **2006**: "Poverty, Political Freedom, and the Roots of Terrorism", *American Economic Review* 96(2): 50-56.
- 2. Abadie, Alberto and Sofia Dermisi, **2006**: "Is terrorism eroding agglomeration economies in central business districts? Lessons from the office real estate market in downtown Chicago". Unpublished manuscript, available online (10/29/2007):

http://ksghome.harvard.edu/~aabadie/chicago.pdf

- 3. Abadie, Alberto and Javier Gardeazabal, **2003**: "The economic costs of conflict: A case study of the Basque country", *American Economic Review* 93(1): 113-132.
- 4. Abadie, Alberto and Javier Gardeazabal, **2007**: "Terrorism and the world economy". Unpublished manuscript, available online (10/29/2007): http://ksghome.harvard.edu/~aabadie/twe.pdf
- 5. Allen, Christopher and Jørgen S. Nielsen, **2002**: "Summary report on islamophobia in the EU after 11 September 2001", report on behalf of the European Monitoring Centre on Racism and Xenophobia, Vienna. Available online (11/25/06):

http://www.eumc.at/eumc/material/pub/anti-islam/Synthesis-report en.pdf

6. American-Arab Anti-Discrimination Committee, **2003**: "Report on hate crimes and discrimination against Arab Americans: the post September 11 backlash, September 11, 2001 - October 11, 2002", Washington D.C.. Available online (12/13/06):

http://www.adc.org/hatecrimes/pdf/2003\_report\_web.pdf

7. Arab American Institute, **2002**: "Healing the nation – The Arab American experience after September 11", Arab American Institute, Washington D.C.. Available online (11/25/06):

http://aai.3cdn.net/64de7330dc475fe470 h1m6b0yk4.pdf

- 8. Åslund, Olof and Dan-Olof Rooth, **2005**: "Shifts in attitudes and labor market discrimination: Swedish experiences after 9-11", *Journal of Population Economics* 18 (4): 603-629.
- 9. Braakmann, Nils, **2007**: "The impact of September 11th, 2001 on the job prospects of foreigners with Arab background Evidence from German labor market data", revised version of University of Lueneburg Working Paper in Economics No. 37, obtained from the author.
- 10. Dávila, Alberto and Marie T. Mora, **2005**: "Changes in the earnings of Arab men in the US between 2000 and 2002", *Journal of Population Economics 18* (4): 587-601.

- 11. European Monitoring Centre on Racism and Xenophobia, **2005**: "The Impact of 7 July 2005 London Bomb Attacks on Muslim Communities in the EU. Available online (11/27/2007): http://www.eumc.at/fra/material/pub/London/London-Bomb-attacks-EN.pdf
- 12. European Monitoring Centre on Racism and Xenophobia, **2006**: "Muslims in the European Union Discrimination and Islamophobia. Available online (11/27/2007): http://fra.europa.eu/fra/material/pub/muslim/Manifestations EN.pdf
- 13. Frey, Bruno S., Simon Luechinger and Alois Stutzer, **2007**: "Calculating tragedy: Assessing the costs of terrorism", *Journal of Economic Surveys* 21(1): 1-24.
- 14. Kaushal, Neeraj, Robert Kaestner and Cordelia Reimers, **2007**: "Labor Market Effects of September 11th on Arab and Muslim Residents of the United States", *The Journal of Human Resources XLII(2)*: 275-308.
- 15. Krueger, Alan B. and Jitka Malečková, **2003**: "Education, poverty, and terrorism: Is there a causal connection?", *Journal of Economic Perspectives* 17(4): 119-144.
- 16. Krueger, Alan B., and David D. Laitin, **2007**: "Kto Kogo?: A cross country study of the origins and targets of terrorism", forthcoming in: Philip Keefer and Norman Loayza (eds.): *Terrorism, Economic Development, and Political Openness*, Cambridge University Press: New York.
- 17. Montalvo, Jose G., **2006**: "Voting after the bombing: Can terrorist attacks change the outcome of democratic elections?". Available at SSRN: http://ssrn.com/abstract=1002833.
- 18. National Commission on Terrorist Attacks Upon the United States, **2004**: "The 9/11 commission report". Available online (11/01/2007): http://www.9-11commission.gov/report/911Report.pdf.
- 19. Orrenius, Pia M. and Madeline Zavodny, **2006**: "Did 9/11 worsen the job prospects of Hispanic immigrants?", Federal Reserve Bank of Dallas Research Department Working Paper 0508.
- 20. Piazza, James A., **2006**: "Rooted in poverty?: Terrorism, poor economic development, and social cleavages", *Terrorism and Political Violence* 18(1): 159-177.
- 21. Sheridan, Lorraine P., **2006**: "Islamophobia Pre- and Post-September 11th, 2001", Journal of Interpersonal Violence 21 (3): 317-336.
- 22. StataCorp, **2005**: "Stata Statistical Software: Release 9.2", StataCorp LP: College Station.

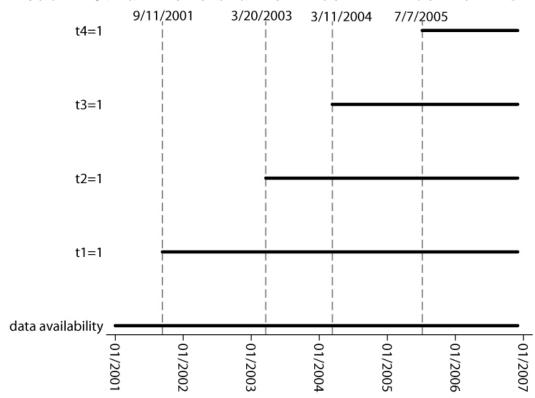
# 8 Appendix

# 8.1 Results and Tables

Table 1: Impact of 9/11 on labor market prospects of Arabs: previous evidence

Study	Country	Treatment / Control groups	Outcome	Results
Dávila, Mora (2005)	US	based on country of birth: US-born white non-Hispanics vs. Middle-East Arab men, vs. African Arab men vs. Iranian, Pakistani, Afghan men	hourly wages	large wage penalty for Middle Eastern Arabs (more than 20%), less for Afghan, Pakistani, Iranian men, no effect for African Arabs
Kaushal, Kaestner, Reimers (2007)	US	US-born (excluding Asians), other immigrants (1st and 2nd generation) vs. 1st and 2nd generation immigrants from variety of "Arab" countries	weekly wages hourly wages hours worked employment intrastate mobility	weekly and hourly wages reduced by approx. 9-11% by 9/11, evidence for temporary decline, no effect on hours worked and employment, intrastate mobility was reduced
Åslund, Rooth (2005)	Sweden	based on country of birth: Middle East, Sweden, other Nordic countries, Western, Eastern Europe Former Yugoslavia, Latin America, Asia, Africa	exits from unemployment	no effect for any group relative to any other
Braakmann (2007)	Germany	based on current and past nationality: Arab countries, Arab countries + Turkey vs. Germans, vs. Central Europeans, vs. South Europeans, vs. East Europeans vs. South-East Europeans, vs. South Americans Americans, vs. Southern Africans	exits from unemployment	no effect for any treatment group compared to any control group

FIGURE 1: OVERVIEW OF OBSERVATION PERIOD AND PERIOD DUMMIES



The lowest black bar marks the observation period ranging from January 2001 to December 2006. The top four bars show the periods where the respective period dummy takes the value "1". The dates of the events of interest are marked by the dashed lines.

Table 2: Descriptive Statistics by group (ethnicity and country of birth)

ced (week)	Arab	British	Other non-white	Arab country	Britain/UK	Other non-European
d (week)						
	0.1728	0.4673	0.2881	0.1585	0.4635	0.3843
	0.3781	0.4989 11.6325	0.4529	0.3653	0.4987	0.4865
	7.0193	7.5445	9069:9	7.4980	7.4885	9.4232
	40.6519	42.8663	41.9650	42.0467	42.8293	42.9256
	13.7213	13.2045	14.0939	13.5715	13.2255	14.0041
Real weekly wage	402.0528	476.0089	438.7444	432.4891	474.1888	532.1755
Age in years	36.3297	280.7903 43.0262	200.9363 37.5025	290.0847 39.5240	42.6776	39.4652
	12.3054	14.0153	12.0349	11.7687	14.0827	11.9794
Married/cohabiting (1 = yes)	0.7197	0.6554	0.4785	0.7374	0.6468	0.6261
N = 2 children made 16	0.4492	0.4752	0.4996	0.4401	0.4780	0.4839
No. of children under 10	1.4900 $1.5418$	0.9079 0.9179	0.0222 $1.0049$	1.5471 $1.5268$	$0.5105 \\ 0.9206$	0.0510
Degree or equivalent ( $1 = yes$ )	0.1096	0.1778	0.1835	0.1262	0.1761	0.2601
11: -1 1 1 1 11	0.3125	0.3824	0.3871	0.3322	0.3809	0.4388
nigher education ( $1 = yes$ )	0.0550 $0.1927$	0.2730	0.0042	0.2067	0.0500	0.0732
GCSE grades $A^*$ -C or equivalent ( $1 = yes$ )	0.1159	0.1823	0.1557	0.0631	0.1853	0.0836
	0.3201	0.3861	0.3626	0.2432	0.3885	0.2768
Other qualification ( $1 = yes$ )	0.2454	0.1235	0.2537	0.3257	0.1224	0.3068
No qualification $(1 = \text{ves})$	0.3684	0.9230	0.1361	0.3520	0.3271	0.4012 $0.1111$
	0.4825	0.3420	0.3429	0.4777	0.3419	0.3143
Health problem hindering at work $(1 = yes)$	0.3185	0.2546	0.1865	0.3452	0.2530	0.1886
Tenure $3$ to $6$ months $(1 \equiv  ext{ves})$	0.4660	0.4356 0.0376	0.3895 0.0521	0.4755	0.4347	0.3912
	0.2192	0.1903	0.2224	0.2076	0.1902	0.2122
Tenure 6 to 12 months $(1 = yes)$	0.0865	0.0686	0.0794	0.0948	0.0687	0.0934
	0.2815	0.2528	0.2705	0.2933	0.2530	0.2910
Tenure 1 to 2 years $(1 = yes)$	0.1635	0.1049	0.1435	0.1469	0.1051	0.1401
Tenure 2 to 5 years $(1 \equiv \text{ves})$	0.3702	0.3065 0.2168	0.3508	0.2891	0.2176	0.3471
( ) C	0.4669	0.4121	0.4422	0.4539	0.4126	0.4413
Tenure 5 to 10 years $(1 = yes)$	0.1394	0.1799	0.1844	0.1635	0.1802	0.1666
•	0.3468	0.3841	0.3880	0.3703	0.3844	0.3727
Tenure 10 to 20 years $(1 = yes)$	0.1466	0.2124	0.1748	0.1422	0.2126	0.1597
T 00 11 01	0.0042	0.4030	0.3300	0.0430	0.4091	40004
tenure more then 20 years $(1 = yes)$	0.0457	0.3552	0.0017	0.2650	0.3537	0.0835 $0.2764$
Firmsize $< 25$ employees $(1 = yes)$	0.2212	0.2468	0.2149	0.2299	0.2468	0.2154
	0.4155	0.4311	0.4109	0.4212	0.4311	0.4112
Firmsize > 500  employees  (1 = yes)	0.1827	0.1572	0.2245	0.1825	0.1571	0.2345
	0.3869	0.3640	0.4174	0.3867	0.3639	0.4238
No. of cases	2,408	130,073	4,328	2,662	132,523	4,905
- Employed	416	60,778	1,247	422	61,424	1,885
- with wage information	401	58,637	1,204	401	59,233	1,800

Means, standard deviation below. Wages have been deflated to 2000 prices using the British Retail Price Index.

Table 3: Descriptive statistics by group (Current nationality and religion)

Variable			Current nationality			Religion	no
	Arab	British	Other n	Other non-European	Muslim	Christian	Other (non Sikh)
	0.1229	0.4595		0.3174	0.1796	0.4573	0.3717
	0.3285	0.4984		0.4656	0.3839	0.4982	0.4834
Hourly real wage	0.0880	11.6467		11.7086	10.3429	11.8275	13.1338
	6.6104	7.5543		9.1653	6.8867	7.6151	8.6605
Actual hours worked (week)	13 8051	42.8330		43.5330	42.0953	42.8345	42.2153
Beal weekly wage	376.4878	15.2405		13.9032	424.2351	13.1307	520.3678
00::	263.5883	287.7845		357.6659	277.1491	289.2781	346.2616
Age in years	34.1675	42.7648		35.2757	37.1218	44.6063	41.0603
	9.6401	14.0116		10.5318	11.9593	13.7541	13.6826
Married/cohabiting (1 = yes)	0.6036	0.6508		0.5509	0.6530	0.6654	0.5564
No. of children under 16	0.4894	0.4767		0.4973	1 2264	0.4715	0.4970
ive of children ander 10	1.2880	0.9408		1.0129	1.4575	0.8881	1.0038
Degree or equivalent ( $1 = yes$ )	0.1031	0.1787		0.1934	0.1245	0.1708	0.2665
	0.3042	0.3831		0.3951	0.3301	0.3763	0.4423
Higher education ( $1 \equiv \mathrm{yes})$	0.0367	0.0802		0.0635	0.0395	0.0842	0.0987
GCSE grades $A^*$ -C or equivalent (1 = yes)	0.0317	0.1816		0.0575	0.0999	0.1759	0.1599
	0.1753	0.3855		0.2329	0.2998	0.3808	0.3666
Other qualification ( $1 = yes$ )	0.4668	0.1266		0.4384	0.2993	0.1323	0.1309
	0.4991	0.3325		0.4963	0.4580	0.3389	0.3374
No qualification $(1 \equiv \text{yes})$	0.2944	0.1383		0.1180	0.3145	0.1354	0.1089
Health problem hindering at work $(1 \equiv \text{ves})$	0.2607	0.2556		0.1248	0.3089	0.2695	0.2610
	0.4392	0.4362		0.3306	0.4621	0.4437	0.4393
Tenure 3 to 6 months $(1 = yes)$	0.1048	0.0378		0.0564	0.0492	0.0350	0.0402
: : : : : : : : : : : : : : : : : : :	0.3076	0.1906		0.2308	0.2164	0.1839	0.1966
tenure o to 12 months $(1 \equiv \text{yes})$	0.1210 0.3274	0.0084		0.1436	0.0803	0.0657	0.0677
Tenure 1 to 2 years $(1 = yes)$	0.2581	0.2929		$0.9969 \\ 0.1946$	0.1246	0.0999	0.1205
	0.4393	0.3064		0.3962	0.3305	0.2999	0.3259
Tenure 2 to 5 years $(1 = yes)$	0.3145	0.2184		0.3141	0.3459	0.2140	0.2326
	0.4662	0.4132		0.4645	0.4761	0.4101	0.4229
tenure 5 to 10 years $(1 = yes)$	0.0045	0.1504		0.1262	0.1799	0.1015	0.2093
Tenure 10 to 20 years $(1 = yes)$	0.0726	0.2125		0.0779	0.1393	0.2147	0.1712
	0.2605	0.4091		0.2681	0.3466	0.4106	0.3771
Tenure more then 20 years $(1 = yes)$	0.0161	0.1458		0.0309	0.0443	0.1599	0.1142
	0.1265	0.3529		0.1731	0.2058	0.3666	0.3183
Firmsize $< 25$ employees $(1 = \mathrm{yes})$	0.2581 $0.4393$	0.2451 $0.4301$		$0.2161 \\ 0.4119$	0.1951 $0.3966$	0.2435 $0.4292$	0.2664 $0.4425$
Firmsize > 500  employees  (1 = yes)	0.1290	0.1600		0.2268	0.2852	0.2019	0.2431
	0.3366	0.3666		0.4191	0.4519	0.4014	0.4294
No. of cases	1,009	139,716		2,347	3,415	81,067	1,276
- Employed	124	64,019		745	610	36,983	473
- with wage information	116	61,716		715	574	35,167	449

Means, standard deviation below. Wages have been deflated to 2000 prices using the British Retail Price Index.

Table 4: Impact of pseudo-interventions

Variable	144			Comptant of hinth	+40441	A	Re	Rollinion
		Ethnicity	Country	y of Dirth	Carrent	nationality		ngion
	vs. British	vs. non-white	vs. British-born	vs. non-European	vs. Britain/UK	vs. non-European	vs. Christians	vs. other religion
Weekly wages:								
Peudo-intervention September 11th	0.0870	0.2753*	0.0421	0.1721	0.4076	0.4748		
Pseudo_intervention war on Iraa	(0.1289)	(0.1366)	(0.1067)	(0.1162)	(0.3801)	(0.4000)	-0.0466	-0 1633
	(0.0695)	(0.0759)	(0.0707)	(0.0797)	(0.1163)	(0.1447)	(0.0652)	(0.0937)
Pseudo-intervention Madrid Bombings	-0.0537	-0.0512	-0.0515	-0.0142	-0.1070	0.0655	-0.0270	0.0686
Pseudo-intervention London Bombings	(0.0725) $0.0810$	(0.0800) 0.0740	(0.0708) 0.0793	(0.0797) 0.0333	(0.1376) 0.1764	(0.1448) 0.0640	(0.0521) 0.0919	(0.0811) -0.0093
Hourly wages:	(0.0703)	(0.0777)	(0.0652)	(0.0722)	(0.1381)	(0.1354)	(0.0514)	(0.0802)
Peudo-intervention September 11th	0.1528	0.2565	0.1139	0.2339*	0.4616	0.4664		
Pseudo-intervention war on Irac	(0.1305) $0.0245$	(0.1360) $-0.0160$	(0.1093) $0.0407$	(0.1169)	(0.3826) $0.0859$	(0.3936)	-0.0684	-0.1884*
	(0.0637)	(0.0720)	(0.0612)	(0.0720)	(0.1044)	(0.1398)	(0.0594)	(0.0865)
Pseudo-intervention Madrid Bombings	-0.0431	-0.0383	-0.0695	-0.0406	-0.0721	-0.0154	-0.0113	0.0618
	(0.0667)	(0.0746)	(0.0623)	(0.0730)	(0.1346)	(0.1491)	(0.0481)	(0.0771)
Fseudo-intervention London Bombings	0.0946	0.1102	0.0914	0.0659	(0.1421	0.0947	0.1309°°°°°°°°°°°°°°°°°°°°°°°°°°°°°°°°°°°°	0.0651
Actual hours worked:						(22.2.2)	(2)22	(22.2.2.)
Peudo-intervention September 11th	-2.4413	1.5086	-1.9263	-0.1631	0.3685	0.7986		
	(1.8914)	(2.4661)	(1.6616)	(2.3095)	(2.9356)	(3.6453)		
Pseudo-intervention war on Iraq	0.4066	1.7603	-0.4863	0.9474	3.2516	4.6472	-0.7052	-2.4059
Peaudo-interrention Madrid Rombings	(1.9537)	(2.3433)	(1.9499)	(2.2041)	(3.6586)	(4.0890)	(1.6723)	(2.7136)
	(1.9375)	(2.4159)	(2.0673)	(2.3571)	(4.4001)	(4.6040)	(1.4612)	(2.6399)
Pseudo-intervention London Bombings	0.8135	0.2061	0.7329	-0.1770	-1.6208	-4.1486	0.7295	-0.6859
Probability of employment:	(1.8208)	(2.2414)	(1.91/2)	(2.1012)	(6077.69)	(1889.6)	(1.4022)	(7.3821)
Peudo-intervention September 11th	0.0628	0.1276	-0.0016	0.0844	0.3679	0.3831		
	(0.1548)	(0.1732)	(0.1422)	(0.1614)	(0.2481)	(0.2797)		
Pseudo-intervention war on Iraq	0.1316	0.2040	0.0115	0.1181	0.1883	0.2315	0.0753	0.2723
	(0.1123)	(0.1329)	(0.1092)	(0.1284)	(0.1810)	(0.2029)	(0.0993)	(0.1640)
Pseudo-intervention Madrid Bombings	0.0125	0.0025	0.0243	-0.0332	-0.4609*	-0.4368	-0.1091	-0.3614*
	(0.1201)	(0.1428)	(0.1151)	(0.1358)	(0.2013)	(0.2319)	(0.0884)	(0.1546)
Pseudo-intervention London Bombings	0.0322	0.0032	0.1485	0.0077	0.5082**	0.2475	0.0921	0.1421
	(0.1097)	(0.1301)	(0.1058)	(0.1219)	(0.1962)	(0.2223)	(0.0812)	(0.1418)

Table 5: Parameters of interest: Impact on real weekly wages, OLS estimates

				Groups defined by:	ed by:			
Variable	Eth	Ethnicity	Count	Country of birth	Current	Current nationality	Re	Religion
	vs. British	vs. non-white	vs. British-born	vs. non-European	vs. Britain/UK	vs. non-European	vs. Christians	vs. other religion
Observed after 9/11	0.0346***	-0.1894*	0.0313***	-0.1170	0.0288**	-0.1747		
	(0.0095)	(0.0753)	(0.0095)	(0.0648)	(0.0094)	(0.1274)		
Observed after 3/20/2003	0.0017	-0.0718	0.0023	0.0425	0.0029	0.0400	0.0133	0.1407
	(0.0110)	(0.0710)	(0.0110)	(0.0602)	(0.0108)	(0.0948)	(0.0141)	(0.0992)
Observed after 3/11/2004	0.0249*	-0.0349	0.0245*	0.0056	0.0264*	-0.0296	0.0258	-0.0004
	(0.0124)	(0.0780)	(0.0122)	(0.0750)	(0.0120)	(0.1105)	(0.0144)	(0.1282)
Observed after $7/7/2005$	0.0102	-0.0069	0.0095	0.0370	0.0115	0.0467	0.0055	0.0849
	(0.0098)	(0.0643)	(0.0098)	(0.0529)	(0.0096)	(0.0898)	(0.0112)	(6680.0)
Treatment group $(1 = yes)$	-0.1938	-0.1406	-0.1831*	-0.2481*	-0.5866	-0.6497	-0.1209*	-0.0061
	(0.1082)	(0.1162)	(0.0903)	(0.0986)	(0.3272)	(0.3405)	(0.0493)	(8690.0)
Post-9/11*Treatment group	0.0515	0.1997	0.0016	0.1201	0.3811	0.4385		
	(0.1178)	(0.1259)	(0.1003)	(0.1091)	(0.3370)	(0.3589)		
Post-Iraq-War*Treatment group	0.0197	-0.0137	0.0515	0.0441	0.1080	0.0756	-0.0421	-0.1596
	(0.0708)	(0.0768)	(0.0724)	(0.0808)	(0.1194)	(0.1482)	(0.0633)	(0.0920)
Post-Madrid-Bombings*Treatment group	-0.0467	-0.0386	-0.0823	-0.0531	-0.1011	0.0594	-0.0381	0.0432
	(0.0738)	(0.0815)	(0.0723)	(0.0804)	(0.1426)	(0.1507)	(0.0531)	(0.0847)
Post-London-Bombings*Treatment group	0.0868	0.0631	0.1187	0.0883	0.1771	0.0838	0.1288*	0.0285
	(0.0702)	(0.0777)	(0.0650)	(0.0718)	(0.1381)	(0.1361)	(0.0513)	(0.0805)
	54.566	1.387	55,121	1,913	57,075	710	32,907	606

Table 6: Parameters of interest: Impact on real hourly wages, OLS estimates

				Groups defined by	ed by:			
Variable	ΕŒ	Ethnicity	Count	Country of birth	Current	Current nationality	Re	Religion
	vs. British	vs. non-white	vs. British-born	vs. non-European	vs. Britain/UK	vs. non-European	vs. Christians	vs. other religion
Observed after 9/11	0.0376***	-0.1942**	0.0347***	-0.1155	0.0321***	-0.1728		
	(0.0089)	(0.0746)	(0.0089)	(0.0640)	(0.0088)	(0.1247)		
Observed after 3/20/2003	0.0075	-0.0363	0.0076	0.0494	0.0086	090'0	0.0075	0.1198
	(0.0100)	(0.0707)	(0.0100)	(0.0609)	(0.008)	(0.0870)	(0.0130)	(0.0956)
Observed after $3/11/2004$	0.0203	-0.0186	0.0200	-0.0267	0.0207	0.0168	0.0186	-0.0388
	(0.0116)	(0.0783)	(0.0114)	(0.0765)	(0.0112)	(0.1194)	(0.0136)	(0.1212)
Observed after 7/7/2005	0.0073	-0.0230	0.0065	0.0201	0.0097	0.0194	0.0018	-0.0155
	(0.0091)	(0.0647)	(0.0091)	(0.0511)	(0.0089)	(0.0843)	(0.0105)	(0.0780)
Treatment group $(1 = yes)$	-0.2251*	-0.1579	-0.1976*	-0.2590*	-0.6054	-0.5848	-0.1209**	-0.0131
	(0.1145)	(0.1206)	(0.0959)	(0.1025)	(0.3278)	(0.3325)	(0.0469)	(0.0659)
Post-9/11*Treatment group	0.0794	0.1687	0.0353	0.1675	0.4370	0.4562		
	(0.1233)	(0.1301)	(0.1052)	(0.1126)	(0.3374)	(0.3514)		
Post-Iraq-War*Treatment group	0.0504	0.0154	0.0812	0.0553	0.0979	0.0745	-0.0442	-0.1664
	(0.0635)	(0.0719)	(0.0613)	(0.0719)	(0.1039)	(0.1401)	(0.0577)	(0.0856)
Post-Madrid-Bombings*Treatment group	-0.0460	-0.0310	-0.1083	-0.0816	-0.0877	-0.0292	-0.0157	0.0368
	(0.0664)	(0.0748)	(0.0629)	(0.0732)	(0.1366)	(0.1526)	(0.0480)	(0.0784)
Post-London-Bombings*Treatment group	0.0863	0.0845	0.1298*	0.1142	0.1426	0.1037	0.1404**	0.0968
	(0.0666)	(0.0739)	(0.0631)	(0.0694)	(0.1486)	(0.1450)	(0.0488)	(0.0731)
7	54.566	1.387	55.121	1.913	57.075	710	32.907	606

21

Table 7: Parameters of interest: Impact on actual weekly working hours, OLS estimates

				Groups defined by:	ted by:			
Variable	Etl	Ethnicity	Count	Country of birth	Current	Current nationality	Re	Religion
	vs. British	vs. non-white	vs. British-born	vs. non-European	vs. Britain/UK	vs. non-European	vs. Christians	vs. other religion
Observed after 9/11	-0.2202	-1.4526	-0.2320	-2.2880	-0.2243	-1.1902		
	(0.2944)	(2.2135)	(0.2936)	(1.8460)	(0.2905)	(2.4232)		
Observed after $3/20/2003$	-0.3958	0.5105	-0.3374	-1.1276	-0.3237	3.1550	0.4338	3.4505
	(0.3484)	(2.6613)	(0.3468)	(1.9713)	(0.3405)	(3.7447)	(0.4328)	(3.6822)
Observed after 3/11/2004	0.1412	-0.2762	$\stackrel{)}{0.0538}$	-2.3643	0.0779	-3.9692	0.2364	0.9071
	(0.3633)	(2.1144)	(0.3598)	(2.1787)	(0.3528)	(5.0521)	(0.4231)	(2.9063)
Observed after $7/7/2005$	0.1684	1.4228	0.1287	0.9700	0.1326	-1.6148	0.0594	3.8568
	(0.2995)	(1.9722)	(0.2983)	(1.4941)	(0.2915)	(2.6338)	(0.3376)	(2.7861)
Treatment group $(1 = yes)$	0.2560	-1.0618	0.1885	-1.7135	-2.1835	-4.1760	0.7884	0.9095
	(1.6254)	(1.9930)	(1.3537)	(1.8188)	(1.5515)	(2.4014)	(1.2195)	(1.8313)
Post- $9/11$ *Treatment group	-0.5250	3.0012	0.0335	2.1067	1.5125	2.5460		
	(2.0534)	(2.5006)	(1.7904)	(2.2747)	(2.8448)	(3.4754)		
Post-Iraq-War*Treatment group	-0.7410	-0.6659	-1.5851	-0.5777	0.5257	0.9300	-1.5409	-3.4803
	(1.9472)	(2.3189)	(1.9568)	(2.2124)	(3.8048)	(4.3146)	(1.6456)	(2.6642)
Post-Madrid-Bombings*Treatment group	-1.6662	-2.3548	-0.1704	-0.5284	0.7195	4.1732	-0.1649	2.3817
	(1.9285)	(2.3616)	(2.0437)	(2.3441)	(4.4980)	(4.6542)	(1.4553)	(2.6076)
Post-London-Bombings*Treatment group	0.9286	0.4723	0.8171	-0.2173	-1.5096	-3.7176	1.1415	-0.2603
	(1.8537)	(2.2679)	(1.9277)	(2.1779)	(3.7764)	(3.8788)	(1.4892)	(2.3857)
	56.532	1,434	57,134	2,000	59.173	735	34.578	957

Table 8: Parameters of interest: Impact on probability of employment, Probit estimates

				Groups defined by	ned by:			
Variable	EFF	Ethnicity	Countr	Country of birth	Current	Current nationality	Re	Religion
	vs. British	vs. non-white	vs. British-born	vs. non-European	vs. Britain/UK	vs. non-European	vs. Christians	vs. other religion
Observed after $9/11$	-0.0505*	-0.0727	-0.0519*	-0.0197	-0.0494*	-0.1455		
	(0.0230)	(0.1249)	(0.0228)	(0.1180)	(0.0223)	(0.1799)		
Observed after 3/20/2003	0.0349	0.0502	0.0415	-0.0111	0.0417	-0.1886	0.0486	0.3821*
	(0.0268)	(0.1407)	(0.0265)	(0.1290)	(0.0259)	(0.2028)	(0.0337)	(0.1933)
Observed after $3/11/2004$	0.0134	-0.0055	0.0142	-0.0523	0.0161	-0.0212	0.0026	0.2831
	(0.0303)	(0.1585)	(0.0299)	(0.1414)	(0.0293)	(0.2188)	(0.0344)	(0.2062)
Observed after $7/7/2005$	-0.0077	-0.0154	-0.0082	-0.1127	92000-	-0.1296	-0.0075	-0.2000
	(0.0234)	(0.1197)	(0.0231)	(0.1070)	(0.0225)	(0.1642)	(0.0262)	(0.1572)
Treatment group $(1 = yes)$	-1.0830***	-0.4157**	-1.0381***	***2897***	-1.5710***	-0.8924***	-0.9412***	-0.3937**
	(0.1232)	(0.1415)	(0.1143)	(0.1293)	(0.2023)	(0.2298)	(0.0725)	(0.1204)
Post-9/11*Treatment group	0.0976	0.1802	-0.0131	0.0926	0.3903	0.4536		
	(0.1413)	(0.1612)	(0.1334)	(0.1512)	(0.2346)	(0.2629)		
Post-Iraq-War*Treatment group	0.0792	0.1274	-0.0161	0.0599	0.0644	0.1087	0.0563	0.1624
	(0.1143)	(0.1346)	(0.1106)	(0.1295)	(0.1826)	(0.2023)	(0.0975)	(0.1610)
Post-Madrid-Bombings*Treatment group	0.0781	0.0861	0.0862	0.0040	-0.3623	-0.3899	-0.0539	-0.2479
	(0.1210)	(0.1436)	(0.1156)	(0.1356)	(0.2058)	(0.2323)	(0.0883)	(0.1531)
Post-London-Bombings*Treatment group	0.0127	-0.0297	0.1374	0.0555	0.5285**	0.3184	0.0623	0.0829
	(0.1102)	(0.1311)	(0.1064)	(0.1226)	(0.1984)	(0.2235)	(0.0817)	(0.1421)
	119.108	5.604	121.435	6.338	126.278	2.702	75,686	3.972

# 8.2 Detailed estimation results

NOT NECESSARILY FOR INCLUSION IN FINAL PAPER, INTERNET APPENDIX OR REFEREE INFORMATION ONLY

Table 9: Weekly real wages, groups defined by ethnicity, OLS-regression

	Compariso	Treatment gr on: British	oup: Ethnicity Pakistan o Comparison: No	or Bangladesh on-White Ethnicity
Variable	naive	adjusted	naive	adjuste
Observed after 9/11	0.0347***	0.0346***	-0.0448	-0.1894
	(0.0083)	(0.0095)	(0.0560)	(0.0753
Observed after $3/20/2003$	0.0052	0.0017	-0.0071	-0.071
Observed after 3/11/2004	$(0.0073) \\ 0.0211**$	(0.0110) 0.0249*	(0.0505) -0.0109	(0.0710 -0.034
Observed after 3/11/2004	(0.0078)	(0.0124)	(0.0567)	(0.0780
Observed after 7/7/2005	0.0191*	0.0102	0.0344	-0.006
	(0.0079)	(0.0098)	(0.0533)	(0.0643
Treatment group $(1 = yes)$	-0.1626	-0.1938	-0.1627	-0.140
D : 0 /44 WT	(0.1170)	(0.1082)	(0.1262)	(0.1162
Post-9/11*Treatment group	-0.0958 (0.1332)	0.0515 $(0.1178)$	-0.0162 (0.1446)	0.199 (0.1259
Post-Iraq-War*Treatment group	0.0397	0.0197	0.0520	-0.013
ose-fraq-war freatment group	(0.0922)	(0.0708)	(0.1051)	(0.0768
Post-Madrid-Bombings*Treatment group	-0.0601	-0.0467	-0.0281	-0.038
0 .	(0.1008)	(0.0738)	(0.1157)	(0.0815
Post-London-Bombings*Treatment group	0.0760	0.0868	0.0607	0.063
	(0.1044)	(0.0702)	(0.1172)	(0.0777
Age in years		0.0809***		0.0610**
A (		(0.0014) -0.0009***		(0.0088
Age (squared)		(0.0000)		-0.0007** (0.0001
Married/cohabiting (1 = yes)		0.0871***		0.041
warred, consisting (1 = yes)		(0.0045)		(0.0272
No. of children under 16		0.0099***		-0.0240
		(0.0021)		(0.0121
Degree or equivalent $(1 = yes)$		0.2189***		0.1640**
		(0.0059)		(0.0389
Higher education (1 = yes)		0.0688***		0.011
CCSE		(0.0065) -0.0459***		(0.0471 -0.048
GCSE grades A*-C or equivalent ( $1 = yes$ )		(0.0053)		(0.0389
Other qualification ( $1 = yes$ )		-0.0731***		-0.050
other quameation ( 1 = jes)		(0.0064)		(0.0359
No qualification $(1 = yes)$		-0.1490***		-0.1372*
		(0.0078)		(0.0457
Tenure 3 to 6 months $(1 = yes)$		-0.0265		0.054
		(0.0159)		(0.0940
Tenure 6 to 12 months $(1 = yes)$		-0.0249		0.026
Tenure 1 to 2 years $(1 = yes)$		(0.0141)		(0.0838
Tenure 1 to 2 years $(1 = yes)$		0.0101 $(0.0134)$		0.106 (0.0814
Tenure 2 to 5 years $(1 = yes)$		0.0553***		0.1562
y (-		(0.0125)		(0.0779
Tenure 5 to 10 years $(1 = yes)$		0.0810***		0.149
		(0.0126)		(0.0840
Fenure 10 to 20 years $(1 = yes)$		0.1144***		0.2631*
		(0.0126)		(0.0847
Tenure more then 20 years $(1 = yes)$		0.1484***		0.2747*
Health problem hindering at work $(1 = ves)$		(0.0130) -0.0685***		(0.0903 -0.1159
rearm problem initidening at work (1 = yes)		(0.0067)		-0.1159 (0.0549
Firmsize $< 25$ employees $(1 = yes)$		-0.0976***		-0.0957*
		(0.0045)		(0.0344
Firmsize $> 500$ employees $(1 = yes)$		0.0752***		0.1008**
		(0.0050)		(0.0291
Constant	5.9607***	4.1352***	5.9608***	4.4830**
	(0.0071)	(0.0578)	(0.0469)	(0.2697
industry fixed effects	no	yes	no	ye
Occupation fixed effects	no	yes	no	ye
Region fixed effects	no	yes	no	ye
Time fixed effects (months, years) N	59,512	yes 54,566	no 1,619	

Table 10: Hourly real wages, groups defined by ethnicity, OLS-regression

	Compariso	on: British	oup: Ethnicity Pakistan o Comparison: No	or Bangladesh on-White Ethnicity
Variable	naive	adjusted	naive	adjuste
Observed after 9/11	0.0410***	0.0376***	-0.0225	-0.1942*
	(0.0080)	(0.0089)	(0.0536)	(0.0746
Observed after 3/20/2003	0.0094	0.0075	-0.0172	-0.036
1 6 9/11/0004	(0.0070)	(0.0100)	(0.0487)	(0.0707
Observed after 3/11/2004	0.0245**	0.0203	0.0367	-0.018
21 1 . 0 7 /7 /0005	(0.0075)	(0.0116)	(0.0516)	(0.0783
Observed after $7/7/2005$	0.0232**	0.0073	-0.0081	-0.023
Treatment group $(1 = \text{yes})$	(0.0075) $-0.2145$	(0.0091) -0.2251*	(0.0488) -0.1964	(0.0647 -0.157
reatment group (1 = yes)	(0.1098)	(0.1145)	(0.1187)	(0.1206
Post-9/11*Treatment group	-0.0139	0.0794	0.0497	0.168
ost-5/11 Treatment group	(0.1249)	(0.1233)	(0.1361)	(0.1301
Oost-Iraq-War*Treatment group	0.0367	0.0504	0.0633	0.015
	(0.0834)	(0.0635)	(0.0965)	(0.0719
Ost-Madrid-Bombings*Treatment group	-0.0245	-0.0460	-0.0367	-0.031
0 0 1	(0.0886)	(0.0664)	(0.1025)	(0.0748
Post-London-Bombings*Treatment group	0.0449	0.0863	0.0762	0.084
	(0.0923)	(0.0666)	(0.1044)	(0.0739
Age in years		0.0586***		0.0346**
		(0.0011)		(0.0081
Age (squared)		-0.0006***		-0.0004**
		(0.0000)		(0.0001
Married/cohabiting (1 = yes)		0.0618***		-0.023
		(0.0042)		(0.0257
No. of children under 16		0.0182***		0.001
		(0.0019)		(0.0112
Degree or equivalent ( $1 = yes$ )		0.2645***		0.1576**
Higher education ( $1 = yes$ )		(0.0057) $0.0902***$		(0.036)
ingher education ( 1 = yes)		(0.0063)		-0.028 (0.0486
GCSE grades A*-C or equivalent ( $1 = yes$ )		-0.0441***		-0.053
GODE grades A -C or equivalent ( 1 = yes)		(0.0048)		(0.0364
Other qualification ( $1 = yes$ )		-0.1114***		-0.0974*
		(0.0059)		(0.0335
No qualification $(1 = yes)$		-0.1767***		-0.1767**
. ,		(0.0070)		(0.0443
Cenure 3 to 6 months (1 = yes)		-0.0105		0.091
, ,		(0.0133)		(0.0847
Cenure 6 to 12 months (1 = yes)		-0.0107		0.031
		(0.0117)		(0.0796
Cenure 1 to 2 years $(1 = yes)$		0.0171		0.093
		(0.0113)		(0.0754)
Cenure 2 to 5 years $(1 = yes)$		0.0462***		0.122
7 (10 (1		(0.0106)		(0.0735
Cenure 5 to 10 years $(1 = yes)$		0.0689***		0.120
7 10 (		(0.0108)		(0.0784
Cenure 10 to 20 years $(1 = yes)$		0.1133*** (0.0107)		0.2409* (0.0804
Cenure more then 20 years (1 = yes)		0.1503***		0.2834*
tenure more then 20 years (1 — yes)		(0.0112)		(0.087)
Health problem hindering at work $(1 = yes)$		-0.0562***		-0.048
reason problem mindering as work (1 = yes)		(0.0061)		(0.0533
Firmsize $< 25$ employees $(1 = yes)$		-0.0999***		-0.0689
(- J)		(0.0041)		(0.0328
Cirmsize > 500  employees  (1 = yes)		0.0855***		0.1150**
		(0.0048)		(0.029)
Constant	2.2348***	0.8874***	2.2166***	1.4594**
	(0.0068)	(0.0491)	(0.0449)	(0.2436
ndustry fixed effects	no	yes	no	ye
Occupation fixed effects	no	yes	no	ye
Region fixed effects	no	yes	no	ye
Time fixed effects (months, years)	no	yes	no	ye

Table 11: Weekly hours worked, groups defined by ethnicity, OLS-regression

	Compariso			Pakistan or Bangladesh arison: Non-White Ethnicity
Variable	naive	adjusted	naive	adjusted
Observed after 9/11	-0.2618	-0.2202	-1.8609	-1.4526
	(0.1849)	(0.2944)	(1.2718)	(2.2135)
Observed after $3/20/2003$	-0.1731	-0.3958	-0.3733	0.5105
01 1 6 0 (44 (000)	(0.1626)	(0.3484)	(1.1981)	(2.6613)
Observed after 3/11/2004	0.2122	0.1412	-0.5229	-0.2762
01 1 . 6 7 /7 /0005	(0.1717)	(0.3633)	(1.3214)	(2.1144)
Observed after $7/7/2005$	0.0772	0.1684	1.6443	1.4228
Treatment group $(1 = yes)$	(0.1652) $0.2692$	(0.2995) $0.2560$	(1.2764) -0.4184	(1.9722) -1.0618
Treatment group (1 – yes)	(1.8905)	(1.6254)	(2.1563)	(1.9930)
Post-9/11*Treatment group	-1.9809	-0.5250	-0.3819	3.0012
1 obt 0/11 11 outment group	(2.3058)	(2.0534)	(2.6327)	(2.5006)
Post-Iraq-War*Treatment group	0.5998	-0.7410	0.7999	-0.6659
	(2.0006)	(1.9472)	(2.3312)	(2.3189)
Post-Madrid-Bombings*Treatment group	-2.7064	-1.6662	-1.9713	-2.3548
	(2.0173)	(1.9285)	(2.4104)	(2.3616)
Post-London-Bombings*Treatment group	1.0729	0.9286	-0.4942	0.4723
	(2.0154)	(1.8537)	(2.3848)	(2.2679)
Age in years		0.7373***		0.7659**
		(0.0375)		(0.2489)
Age (squared)		-0.0090***		-0.0086**
25 1 1/ 1 111 /5		(0.0005)		(0.0031)
Married/cohabiting (1 = yes)		0.9464***		1.7168
N f . 1 '11 1 16		(0.1423)		(0.8760)
No. of children under 16		-0.0357 (0.0646)		-0.9763** (0.3327)
Degree or equivalent ( $1 = ves$ )		-0.6388***		-0.2919
Degree of equivalent (1 — yes)		(0.1701)		(1.1683)
Higher education ( $1 = ves$ )		-0.3993*		0.2420
ingher education (1 = yes)		(0.2017)		(1.4548)
GCSE grades $A^*$ -C or equivalent (1 = yes)		0.0000		0.4571
		(0.1568)		(1.1846)
Other qualification ( $1 = yes$ )		1.3199***		1.2059
		(0.1983)		(1.1053)
No qualification $(1 = yes)$		0.3761		0.1970
		(0.2481)		(1.6177)
Tenure 3 to 6 months $(1 = yes)$		-0.2568		-1.7219
		(0.4030)		(2.1414)
Tenure 6 to 12 months $(1 = yes)$		-0.5876		-1.5263
T (1 )		(0.3637)		(2.1950)
Tenure 1 to 2 years $(1 = yes)$		-0.5075		-0.0044
T 2 to 5 (1)		(0.3379)		(1.8629)
Tenure 2 to 5 years $(1 = yes)$		-0.1657		0.4174 $(1.8468)$
Tenure 5 to 10 years $(1 = yes)$		(0.3154) -0.2630		-0.1842
Tenure 5 to 10 years (1 — yes)		(0.3219)		(2.0083)
Tenure 10 to 20 years $(1 = yes)$		-0.6982*		-1.1452
Tenare 10 to 20 years (1 = yes)		(0.3210)		(2.0581)
Tenure more then 20 years $(1 = yes)$		-0.9174**		-0.2744
3		(0.3355)		(2.2351)
Health problem hindering at work $(1 = yes)$		-0.7331***		0.1725
		(0.1960)		(1.5411)
Firmsize $< 25$ employees $(1 = yes)$		0.5108***		-0.7623
		(0.1306)		(0.9163)
Firmsize $> 500$ employees (1 = yes)		-0.6819***		1.0460
_		(0.1551)		(0.9625)
Constant	43.0880***	23.2850***	43.7756***	20.5268***
	(0.1564)	(1.0974)	(1.0390)	(5.9339)
Industry fixed effects	no	yes	no	yes
Occupation fixed effects	no	yes	no	yes
Region fixed effects	no	yes	no	yes
Time fixed effects (months, years)  N	61,668	yes 56,532	1,677	yes 1,434
14	01,008	50,552	1,077	1,434

Table 12: Probability of employment, groups defined by ethnicity, Probit-regression

1 ROBIT-REGRESSION						
	Treatment group: Ethnicity Pakistan or Bangladesh					
	Compariso	Comparison: British		Comparison: Non-White Ethnicity		
Variable	naive	adjusted	naive	adjusted		
Observed after 9/11	-0.0393**	-0.0505*	-0.0917	-0.0727		
	(0.0124)	(0.0230)	(0.0744)	(0.1249)		
Observed after 3/20/2003	-0.0254*	0.0349	-0.0143	0.0502		
	(0.0106)	(0.0268)	(0.0617)	(0.1407)		
Observed after 3/11/2004	-0.0109	0.0134	-0.0032	-0.0055		
	(0.0113)	(0.0303)	(0.0659)	(0.1585)		
Observed after 7/7/2005	-0.0610***	-0.0077	-0.0265	-0.0154		
	(0.0106)	(0.0234)	(0.0611)	(0.1197)		
Treatment group $(1 = yes)$	-1.0204***	-1.0830***	-0.5741***	-0.4157**		
	(0.1011)	(0.1232)	(0.1196)	(0.1415)		
Post-9/11*Treatment group	0.0916	0.0976	0.1440	0.1802		
	(0.1156)	(0.1413)	(0.1369)	(0.1612)		
Post-Iraq-War*Treatment group	0.0853	0.0792	0.0742	0.1274		
	(0.0922)	(0.1143)	(0.1104)	(0.1346)		
Post-Madrid-Bombings*Treatment group	0.0348	0.0781	0.0271	0.0861		
	(0.0974)	(0.1210)	(0.1171)	(0.1436)		
Post-London-Bombings*Treatment group	0.0535	0.0127	0.0189	-0.0297		
0 .	(0.0902)	(0.1102)	(0.1084)	(0.1311)		
Age in years	, ,	0.1862***	` /	0.1081***		
		(0.0024)		(0.0128)		
Age (squared)		-0.0026***		-0.0014***		
J ( 1 )		(0.0000)		(0.0002)		
Married/cohabiting (1 = ves)		0.5755***		0.4236***		
,		(0.0111)		(0.0542)		
No. of children under 16		-0.0783***		-0.0617**		
		(0.0051)		(0.0203)		
Degree or equivalent ( $1 = \text{yes}$ )		0.0009		0.1660*		
3		(0.0140)		(0.0716)		
Higher education ( $1 = yes$ )		0.0135		0.1200		
( , , , ,		(0.0162)		(0.0876)		
GCSE grades A*-C or equivalent ( $1 = yes$ )		-0.0746***		-0.0491		
8 ( - ))		(0.0124)		(0.0696)		
Other qualification ( $1 = yes$ )		-0.1311***		-0.2867***		
( , , , ,		(0.0148)		(0.0641)		
No qualification $(1 = yes)$		-0.4706***		-0.6502***		
1		(0.0160)		(0.0802)		
Health problem hindering at work $(1 = yes)$		-0.9988***		-0.9138***		
3 - 7		(0.0112)		(0.0668)		
Constant	-0.0096	-2.8425***	-0.4558***	-2.8042***		
	(0.0106)	(0.0995)	(0.0648)	(0.4154)		
Occupation fixed effects	no	yes	no	yes		
Region fixed effects	no	yes	no	yes		
Time fixed effects (months, years)	no	yes	no	yes		
N	132,955	119,108	6,750	5,604		
-11	102,000	110,100	0,730	5,004		

Table 13: Weekly real wages, groups defined by country of birth, OLS-regression

	Treatment group: Comparison: British-born		Born in Arabian / Muslim country Comparison: Non-European-born	
Variable	naive	adjusted	naive	adjusted
Observed after 9/11	0.0312***	0.0313***	0.0026	-0.1170
	(0.0083)	(0.0095)	(0.0577)	(0.0648
Observed after $3/20/2003$	0.0066	0.0023	-0.0472	0.042
	(0.0073)	(0.0110)	(0.0447)	(0.0602
Observed after 3/11/2004	0.0194*	0.0245*	0.0300	0.0056
	(0.0078)	(0.0122)	(0.0488)	(0.0750
Observed after $7/7/2005$	0.0215**	0.0095	-0.0647	0.0370
	(0.0078)	(0.0098)	(0.0451)	(0.0529
Treatment group $(1 = yes)$	-0.0705	-0.1831*	-0.2228	-0.2481
	(0.1022)	(0.0903)	(0.1145)	(0.0986
Post-9/11*Treatment group	-0.0832	0.0016	-0.0545	0.120
D . I III * M	(0.1206)	(0.1003)	(0.1337)	(0.1091
Post-Iraq-War*Treatment group	-0.0183	0.0515	0.0355	0.044
D M. 1 . 1 D. 1	(0.0924)	(0.0724)	(0.1026)	(0.0808
Post-Madrid-Bombings*Treatment group	0.0066	-0.0823	-0.0040	-0.053
Post-London-Bombings*Treatment group	$(0.0986) \\ 0.0555$	$(0.0723) \\ 0.1187$	(0.1099) $0.1417$	(0.0804 0.088
rost-London-Bombings Treatment group	(0.0998)	(0.0650)	(0.1094)	(0.0718
Age in years	(0.0996)	0.0816***	(0.1094)	0.0603**
Age in years		(0.0014)		(0.0076
Age (squared)		-0.0009***		-0.0007**
Age (squared)		(0.0000)		(0.0001
Married/cohabiting (1 - ves)		0.0892***		0.0556
Married/cohabiting (1 = yes)		(0.0045)		(0.0243
No. of children under 16		0.0080***		0.001
		(0.0021)		(0.0112
Degree or equivalent ( $1 = yes$ )		0.2222***		0.1806**
		(0.0059)		(0.0337
$ \   \text{Higher education (1 = yes)} $		0.0680***		0.067
		(0.0065)		(0.0406
GCSE grades A*-C or equivalent ( $1 = yes$ )		-0.0435***		-0.004
GODE grades if -o or equivalent ( 1 = yes)		(0.0053)		(0.0380
Other qualification ( $1 = yes$ )		-0.0756***		0.034
(- <i>y</i> )		(0.0064)		(0.0317
No qualification $(1 = yes)$		-0.1482***		-0.073
1		(0.0078)		(0.0459
Tenure 3 to 6 months $(1 = yes)$		-0.0199		0.033
		(0.0159)		(0.0732
Tenure 6 to 12 months $(1 = yes)$		-0.0210		0.009
		(0.0141)		(0.0618
Tenure 1 to 2 years $(1 = yes)$		0.0098		0.069
3		(0.0134)		(0.0594
Tenure 2 to 5 years $(1 = yes)$		0.0578***		0.1189
		(0.0125)		(0.0560
Tenure 5 to 10 years $(1 = yes)$		0.0856***		0.073
Tenure 10 to 20 years $(1 = yes)$		(0.0126)		(0.0595
		0.1181***		0.2042**
		(0.0126)		(0.0605
Tenure more then 20 years $(1 = yes)$		0.1532***		0.2687**
		(0.0131)		(0.0652
Health problem hindering at work $(1 = yes)$		-0.0679***		-0.0935
nearth problem mindering at work (1 = yes)		(0.0067)		(0.0418
Firmsize $< 25$ employees $(1 = yes)$		-0.0967***		-0.1690**
\ / /		(0.0045)		(0.0290
Firmsize $> 500$ employees (1 = yes)		0.0733***		0.1009**
1 13 11 ( 3 11)		(0.0050)		(0.0258
Constant	5.9590***	4.1247***	6.1113***	4.4670**
	(0.0071)	(0.0573)	(0.0515)	(0.1841
industry fixed effects	no	yes	no	ye
Occupation fixed effects	no	yes	no	ye
Region fixed effects	no	yes	no	ye
Time fixed effects (months, years)	no	yes	no	ye
N	60,116	55,121	2,211	1,91

Table 14: Hourly real wages, groups defined by country of birth,  ${\rm OLS\textsc{-}regression}$ 

	Treatment group: Comparison: British-born		Born in Arabian / Muslim country Comparison: Non-European-born	
Variable	naive	adjusted	naive	adjusted
Observed after 9/11	0.0390***	0.0347***	-0.0199	-0.1155
	(0.0079)	(0.0089)	(0.0538)	(0.0640)
Observed after 3/20/2003	0.0104	0.0076	-0.0407	0.049
	(0.0069)	(0.0100)	(0.0454)	(0.0609
Observed after 3/11/2004	0.0231**	0.0200	0.0522	-0.026
01 1 6 7 7 000 7	(0.0074)	(0.0114)	(0.0482)	(0.0765
Observed after $7/7/2005$	0.0250***	0.0065	-0.0766	0.020
TD (1 )	(0.0074)	(0.0091)	(0.0438)	(0.0511
Treatment group $(1 = yes)$	-0.0908	-0.1976*	-0.2615*	-0.2590
Post-9/11*Treatment group	(0.1020) -0.0246	(0.0959) $0.0353$	(0.1122) 0.0343	$(0.1025 \\ 0.167$
rost-9/11 Treatment group	(0.1184)	(0.1052)	(0.1300)	(0.1126
Post-Iraq-War*Treatment group	-0.0432	0.0812	0.0079	0.055
rost-fraq-war freatment group	(0.0861)	(0.0613)	(0.0973)	(0.0719
Post-Madrid-Bombings*Treatment group	-0.0038	-0.1083	-0.0329	-0.081
	(0.0912)	(0.0629)	(0.1031)	(0.0732
Post-London-Bombings*Treatment group	0.0728	0.1298*	0.1744	0.1142
3 1	(0.0941)	(0.0631)	(0.1037)	(0.0694
Age in years	(	0.0588***	(,	0.0431***
		(0.0011)		(0.0075
Age (squared)		-0.0006***		-0.0005***
		(0.0000)		(0.0001
Married/cohabiting (1 = yes)		0.0631***		0.016
		(0.0042)		(0.0239
No. of children under 16  Degree or equivalent ( $1 = yes$ )		0.0166***		0.016
		(0.0019)		(0.0109
		0.2669***		0.1817**
Higher education (1 = yes)		(0.0057)		(0.0312
		0.0885***		0.037
CCCP 1 A*C : 1 · / · · · )		(0.0063)		(0.0406
GCSE grades A*-C or equivalent ( $1 = yes$ )		-0.0430***		0.012
Otherlifetime ( 1)		(0.0047) -0.1143***		(0.0364
Other qualification ( $1 = yes$ )				0.009 (0.0300
No qualification $(1 = yes)$		(0.0058) -0.1759***		-0.1047
No quantication (1 = yes)		(0.0069)		(0.0434
Tenure 3 to 6 months $(1 = yes)$		-0.0043		0.032
renure 5 to 6 months (1 = yes)		(0.0133)		(0.0642
Tenure 6 to 12 months $(1 = \text{ves})$		-0.0052		-0.0003
rendre o to 12 months (1 = yes)		(0.0117)		(0.0557
Tenure 1 to 2 years $(1 = yes)$		0.0214		0.062
<del>-</del> y (- y)		(0.0113)		(0.0538
Tenure 2 to 5 years $(1 = yes)$		0.0516***		0.083
		(0.0106)		(0.0495
Tenure 5 to 10 years $(1 = yes)$		0.0764***		0.047
		(0.0108)		(0.0535
Tenure 10 to 20 years $(1 = yes)$		0.1205***		0.1882***
Tenure more then 20 years (1 = yes)		(0.0108)		(0.0550
		0.1587***		0.2457**
		(0.0113)		(0.0609
Health problem hindering at work $(1 = yes)$		-0.0552***		-0.052
		(0.0061)		(0.0390
Firmsize < 25 employees (1 = yes)		-0.0995***		-0.1402**
7' ' > 500 ' 1 ' '		(0.0041)		(0.0273
Firmsize $> 500$ employees (1 = yes)		0.0833***		0.1045**
g	0.0000***	(0.0047)	0.4005***	(0.0260
Constant	2.2328***	0.8828***	2.4035***	1.1150**
- Jt C Jt	(0.0068)	(0.0487)	(0.0467)	(0.2302
Industry fixed effects	no	yes	no	ye
Occupation fixed effects Region fixed effects	no no	yes	no no	ye
Region fixed effects  Time fixed effects (months, years)	no no	yes yes	no no	ye
N	60,116	55,121	2,211	ye 1,91

Table 15: Weekly hours worked, groups defined by country of birth,  $$\operatorname{OLS}-\operatorname{regression}$$ 

	Comparison:		p: Born in Arabian / Muslim country Comparison: Non-European-born	
Variable	naive	adjusted	naive	adjusted
Observed after 9/11	-0.2841	-0.2320	-1.0539	-2.2880
	(0.1845)	(0.2936)	(1.2387)	(1.8460)
Observed after $3/20/2003$	-0.1168	-0.3374	-0.5968	-1.1276
	(0.1624)	(0.3468)	(0.9983)	(1.9713)
Observed after 3/11/2004	0.1047	0.0538	-0.0405	-2.3643
01 1 . 6 . 7 /7 /0005	(0.1712)	(0.3598)	(1.0574)	(2.1787)
Observed after $7/7/2005$	0.1346	0.1287	0.7077	0.9700
Treatment group $(1 = yes)$	(0.1642) $0.5879$	(0.2983)	(0.9520) -0.4380	(1.4941) -1.7135
Treatment group $(1 = yes)$	(1.5237)	0.1885 $(1.3537)$	(1.8682)	(1.8188)
Post-9/11*Treatment group	-1.0723	0.0335	-0.3025	2.1067
1 050-5/11 Treatment group	(1.9572)	(1.7904)	(2.3123)	(2.2747)
Post-Iraq-War*Treatment group	0.0482	-1.5851	0.5282	-0.5777
1 ook mad War Trouvment Stoap	(1.9827)	(1.9568)	(2.2176)	(2.2124)
Post-Madrid-Bombings*Treatment group	-0.6167	-0.1704	-0.4714	-0.5284
3	(2.1122)	(2.0437)	(2.3599)	(2.3441)
Post-London-Bombings*Treatment group	-0.2967	0.8171	-0.8697	-0.2173
- ·	(2.0770)	(1.9277)	(2.2827)	(2.1779)
Age in years	. ,	0.7353***	, ,	0.4550*
		(0.0374)		(0.2267)
Age (squared)		-0.0089***		-0.0055*
		(0.0005)		(0.0028)
Married/cohabiting (1 = yes)		1.0219***		1.4501
		(0.1413)		(0.8123)
No. of children under 16		-0.0649		-0.6924*
Degree or equivalent ( $1 = yes$ )		(0.0642)		(0.3149)
		-0.6380***		1.2408
TT: 1 1 1: (1 )		(0.1689)		(1.0440)
Higher education ( $1 = yes$ )		-0.4528*		0.9901
GCSE grades A*-C or equivalent ( $1 = ves$ )		$(0.2006) \\ 0.0605$		(1.2264) -0.4519
GCSE grades A -C or equivalent (1 = yes)		(0.1558)		(1.2333)
Other qualification ( $1 = yes$ )		1.2170***		1.4563
Other qualification ( 1 = yes)		(0.1975)		(0.9881)
No qualification $(1 = yes)$		0.3000		2.8559*
<b>1</b> (-		(0.2467)		(1.4198)
Tenure 3 to 6 months $(1 = yes)$		-0.2551		2.2229
		(0.4031)		(1.5574)
Tenure 6 to 12 months $(1 = yes)$		-0.6447		1.9343
		(0.3624)		(1.6949)
Tenure 1 to 2 years $(1 = yes)$		-0.6492		1.4813
		(0.3360)		(1.4111)
Tenure 2 to 5 years $(1 = yes)$		-0.3094		2.1350
		(0.3140)		(1.3250)
Tenure 5 to 10 years $(1 = yes)$		-0.3258		1.7476
T		(0.3203)		(1.4239)
Tenure 10 to 20 years $(1 = yes)$		-0.7602*		0.3993
The second of th		(0.3193)		(1.5057)
Tenure more then 20 years $(1 = yes)$		-1.0540** (0.3342)		3.0405 (1.6281)
Health problem hindering at work $(1 = yes)$		-0.7262***		(1.6281) -1.6208
Health problem hindering at work $(1 = yes)$		(0.1959)		(1.0492)
Firmsize $< 25$ employees $(1 = yes)$		0.5077***		-0.5130
1 minute ( 20 cmployees (1 — yes)		(0.1298)		(0.7909)
Firmsize $> 500$ employees (1 = yes)		-0.6128***		0.5789
		(0.1548)		(0.7923)
Constant	43.0706***	23.4777***	44.0965***	32.1829***
	(0.1561)	(1.0963)	(1.0878)	(5.2850)
Industry fixed effects	no	yes	no	yes
Occupation fixed effects	no	yes	no	yes
Region fixed effects	no	yes	no	yes
Time fixed effects (months, years)	no	yes	no	yes

Table 16: Probability of employment, groups defined by country of birth, Probit-regression

		Treatment group British-born	e: Born in Arabian / Muslim country Comparison: Non-European-born	
Variable	naive	adjusted	naive	adjusted
Observed after 9/11	-0.0436***	-0.0519*	-0.1344	-0.0197
,	(0.0123)	(0.0228)	(0.0695)	(0.1180)
Observed after 3/20/2003	-0.0215*	0.0415	-0.0184	-0.0111
, ,	(0.0105)	(0.0265)	(0.0568)	(0.1290)
Observed after 3/11/2004	-0.0164	0.0142	0.0249	-0.0523
	(0.0111)	(0.0299)	(0.0595)	(0.1414)
Observed after 7/7/2005	-0.0578***	-0.0082	0.0047	-0.1127
	(0.0105)	(0.0231)	(0.0531)	(0.1070)
Treatment group $(1 = ves)$	-0.9326***	-1.0381***	-0.7772***	-0.5897***
	(0.0964)	(0.1143)	(0.1133)	(0.1293)
Post-9/11*Treatment group	-0.0525	-0.0131	0.0383	0.0926
1 050-5/11 Treatment group	(0.1112)	(0.1334)	(0.1306)	(0.1512)
Post-Iraq-War*Treatment group	0.0483	-0.0161	0.0451	0.0599
1 ost-maq- war Treatment group	(0.0886)	(0.1106)	(0.1047)	(0.1295)
Post-Madrid-Bombings*Treatment group	0.0161	0.0862	-0.0252	0.0040
1 ost-wadrid-bombings Treatment group	(0.0942)	(0.1156)	(0.1109)	(0.1356)
Post-London-Bombings*Treatment group	0.1548	0.1374	0.0923	0.0555
	(0.0885)	(0.1064)	(0.1027)	(0.1226)
Age in years	(0.0883)	0.1875***	(0.1027)	0.1132***
Age in years		(0.0023)		(0.0124)
Age (squared)		-0.0026***		-0.0015***
Age (squared)		(0.0000)		(0.0001)
Married/cohabiting (1 = yes)		0.5828***		0.4627***
Married/conabiting (1 = yes)		(0.0110)		
No. of children under 16		-0.0818***		(0.0506) -0.0481**
No. of children under 16				
5		(0.0051)		(0.0184)
Degree or equivalent ( $1 = yes$ )		0.0078		0.1031
TT: 1 1 1 / 1 / 1		(0.0139)		(0.0625)
Higher education ( $1 = yes$ )		0.0093		0.0589
0000		(0.0161)		(0.0836)
GCSE grades A*-C or equivalent ( $1 = yes$ )		-0.0706***		-0.0207
		(0.0122)		(0.0805)
Other qualification ( $1 = yes$ )		-0.1403***		-0.3210***
		(0.0147)		(0.0590)
No qualification $(1 = yes)$		-0.4714***		-0.6872***
		(0.0159)		(0.0779)
Health problem hindering at work $(1 = yes)$		-0.9936***		-0.8303***
		(0.0111)		(0.0554)
Constant	-0.0155	-2.3424***	-0.1708**	-2.9253***
	(0.0106)	(0.0770)	(0.0604)	(0.4807)
Occupation fixed effects	no	yes	no	yes
Region fixed effects	no	yes	no	yes
Time fixed effects (months, years)	no	yes	no	yes
N	135,667	121,435	7,577	6,338

Table 17: Weekly real wages, groups defined by current nationality, OLS-regression

		ient group: Ci on: British	urrent nationality from Arab / Muslim country Comparison: Non-European		
Variable	naive	adjusted	naive	adjusted	
Observed after 9/11	0.0309***	0.0288**	-0.1234	-0.174	
	(0.0081)	(0.0094)	(0.0910)	(0.1274	
Observed after $3/20/2003$	0.0050	0.0029	-0.0336	0.0400	
01 0 0 144 (0004	(0.0071)	(0.0108)	(0.0724)	(0.0948)	
Observed after 3/11/2004	0.0201**	0.0264*	-0.0229	-0.0296	
Observed after 7/7/2005	$(0.0077) \\ 0.0208**$	$(0.0120) \\ 0.0115$	(0.0780) -0.0550	(0.1105 0.046'	
Observed after 1/1/2005	(0.0077)	(0.0096)	(0.0728)	(0.0898)	
Treatment group $(1 = yes)$	-0.4733	-0.5866	-0.6464*	-0.649	
readment group (1 — yes)	(0.3052)	(0.3272)	(0.3166)	(0.3405	
Post-9/11*Treatment group	0.1799	0.3811	0.3342	0.438	
, , , , , , , , , , , , , , , , , , , ,	(0.3310)	(0.3370)	(0.3451)	(0.3589	
Post-Iraq-War*Treatment group	-0.0481	0.1080	-0.0095	0.075	
	(0.1640)	(0.1194)	(0.1801)	(0.1482	
Post-Madrid-Bombings*Treatment group	-0.0767	-0.1011	-0.0337	0.059	
D . T . D . L	(0.1536)	(0.1426)	(0.1729)	(0.1507	
Post-London-Bombings*Treatment group	0.2573	0.1771	0.3331	0.0838	
A	(0.1584)	(0.1381)	(0.1750)	(0.1361	
Age in years		0.0812*** (0.0013)		0.0412** (0.0128	
Age (squared)		-0.0009***		-0.0005**	
Age (squared)		(0.0000)		(0.0003	
Married/cohabiting (1 = yes)		0.0877***		0.022	
married, conditions (1 = yes)		(0.0044)		(0.0406)	
No. of children under 16		0.0073***		-0.0168	
		(0.0021)		(0.0213	
Degree or equivalent ( $1 = yes$ )		0.2191***		0.1186	
		(0.0058)		(0.0658)	
Higher education ( $1 = yes$ )		0.0647***		0.0548	
		(0.0064)		(0.0923)	
GCSE grades A*-C or equivalent ( $1 = yes$ )		-0.0427***		-0.0816	
Other qualification ( $1 = yes$ )		(0.0052) -0.0725***		(0.0867)	
Other quantication ( $1 = yes$ )		(0.0063)		0.0304 (0.0566)	
No qualification $(1 = yes)$		-0.1483***		0.000	
(1 = yes)		(0.0076)		(0.0877	
Tenure 3 to 6 months $(1 = yes)$		-0.0161		0.0073	
3,		(0.0156)		(0.1046)	
Tenure 6 to 12 months $(1 = yes)$		-0.0192		0.0518	
		(0.0139)		(0.0864)	
Tenure 1 to 2 years $(1 = yes)$		0.0124		0.1342	
		(0.0132)		(0.0804)	
Tenure 2 to 5 years $(1 = yes)$		0.0596***		0.1738	
T (10 (1 )		(0.0123)		(0.0804)	
Tenure 5 to 10 years $(1 = yes)$		0.0839***		0.1248	
T 10 t- 20 (1)		(0.0124) 0.1185***		(0.0959) 0.4314***	
Tenure 10 to 20 years $(1 = yes)$		(0.0124)		(0.1059)	
Tenure more then 20 years $(1 = yes)$		0.1534***		0.4116**	
renare more then 20 years (1 = yes)		(0.0129)		(0.1460)	
Health problem hindering at work $(1 = yes)$		-0.0688***		-0.0547	
meanth problem findering at work (1 – yes)		(0.0066)		(0.1113	
Firmsize $< 25$ employees $(1 = yes)$		-0.0986***		-0.1193	
· ( = (-		(0.0044)		(0.0501	
Firmsize $> 500$ employees (1 = yes)		0.0751***		0.1165	
		(0.0049)		(0.0488	
Constant	5.9639***	3.8519***	6.1370***	5.1782***	
	(0.0069)	(0.0379)	(0.0777)	(0.2779	
Industry fixed effects	no	yes	no	ye	
Occupation fixed effects	no	yes	no	ye	
Region fixed effects	no	yes	no	yea	
Time fixed effects (months, years) N	62,326	yes 57,075	no 835	yes 710	

Table 18: Hourly real wages, groups defined by current nationality, OLS-regression

		ent group: Cu on: British	urrent nationality from Arab / Muslim country Comparison: Non-European		
Variable	naive	adjusted	naive	adjusted	
Observed after 9/11	0.0384***	0.0321***	-0.1294	-0.1728	
	(0.0078)	(0.0088)	(0.0866)	(0.1247	
Observed after 3/20/2003	0.0091	0.0086	-0.0442	0.060'	
01 0 0 144 (0004	(0.0068)	(0.0098)	(0.0760)	(0.0870)	
Observed after 3/11/2004	0.0232**	0.0207	0.0659	0.0168	
Observed after 7/7/2005	(0.0073) $0.0250***$	(0.0112) $0.0097$	(0.0775) -0.0902	(0.1194 0.019	
Observed after 1/1/2005	(0.0073)	(0.0089)	(0.0671)	(0.0843)	
Treatment group $(1 = yes)$	-0.4574	-0.6054	-0.6019	-0.5848	
readment group (1 — yes)	(0.2977)	(0.3278)	(0.3080)	(0.3325	
Post-9/11*Treatment group	0.1919	0.4370	0.3597	0.456	
, , , , , , , , , , , , , , , , , , , ,	(0.3211)	(0.3374)	(0.3343)	(0.3514	
Post-Iraq-War*Treatment group	-0.1038	0.0979	-0.0505	0.074	
	(0.1490)	(0.1039)	(0.1679)	(0.1401	
Post-Madrid-Bombings*Treatment group	-0.0697	-0.0877	-0.1124	-0.0295	
D . T . 1 . D . 1: * W	(0.1615)	(0.1366)	(0.1799)	(0.1526)	
Post-London-Bombings*Treatment group	0.2275	0.1426	0.3427	0.1037	
A :	(0.1815)	(0.1486) $0.0587***$	(0.1943)	(0.1450 0.0299	
Age in years		(0.0011)		(0.0121	
Age (squared)		-0.0006***		-0.0003	
rige (squared)		(0.0000)		(0.0002	
Married/cohabiting (1 = yes)		0.0614***		0.0068	
		(0.0041)		(0.0395)	
No. of children under 16		0.0161***		-0.0164	
		(0.0019)		(0.0211	
Degree or equivalent ( $1 = yes$ )		0.2627***		0.1148	
		(0.0056)		(0.0620)	
Higher education ( $1 = yes$ )		0.0849***		0.000	
GGGD 1 A*G : 1 (/1 )		(0.0062)		(0.0925)	
GCSE grades A*-C or equivalent ( $1 = yes$ )		-0.0419***		-0.0689	
Other qualification ( $1 = yes$ )		(0.0047) -0.1108***		(0.0816) 0.0042	
Other quantication (1 = yes)		(0.0057)		(0.0527	
No qualification $(1 = yes)$		-0.1756***		-0.0413	
(- //		(0.0068)		(0.0863)	
Tenure 3 to 6 months $(1 = yes)$		-0.0012		0.0241	
, ,		(0.0131)		(0.0964)	
Tenure 6 to 12 months $(1 = yes)$		-0.0043		0.0530	
		(0.0116)		(0.0798)	
Tenure 1 to 2 years $(1 = yes)$		0.0227*		0.1306	
		(0.0111)		(0.0785)	
Tenure 2 to 5 years $(1 = yes)$		0.0526***		0.1264	
T. (10 (1)		(0.0105)		(0.0761)	
Tenure 5 to 10 years $(1 = yes)$		0.0738***		0.0812 (0.0909)	
Tenure 10 to 20 years $(1 = yes)$		(0.0107) $0.1190***$		0.4108***	
Tenure 10 to 20 years (1 — yes)		(0.0106)		(0.1017)	
Tenure more then 20 years $(1 = yes)$		0.1574***		0.3686	
remare more than 20 years (1 = yes)		(0.0111)		(0.1440)	
Health problem hindering at work $(1 = yes)$		-0.0551***		-0.0260	
- 0 ( 3 %)		(0.0059)		(0.1014	
Firmsize < 25  employees  (1 = yes)		-0.1015***		-0.055	
1 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1		(0.0040)		(0.0508	
Firmsize $> 500$ employees $(1 = yes)$		0.0847***		0.1403*	
_		(0.0047)	white	(0.0474	
Constant	2.2380***	0.5774***	2.3824***	1.8376***	
	(0.0067)	(0.0306)	(0.0719)	(0.2651	
Industry fixed effects	no	yes	no	ye	
Occupation fixed effects	no	yes	no	ye	
Region fixed effects Time fixed effects (months, years)	no no	yes yes	no no	yes	
N N	62,326	57,075	835	yes 710	

Table 19: Weekly hours worked, groups defined by current nationality, OLS-regression

	Treatment group: Curre Comparison: British		ent nationality from Arab / Muslim country Comparison: Non-European	
Variable	naive	adjusted	naive	adjusted
Observed after 9/11	-0.2883	-0.2243	-1.2268	-1.1902
	(0.1813)	(0.2905)	(1.7221)	(2.4232)
Observed after 3/20/2003	-0.1208	-0.3237	-0.3034	3.1550
	(0.1592)	(0.3405)	(1.6485)	(3.7447)
Observed after 3/11/2004	0.1315	0.0779	-1.8890	-3.9692
01 1 0	(0.1677)	(0.3528)	(1.8439)	(5.0521)
Observed after $7/7/2005$	0.1248	0.1326	0.8310	-1.6148
Treatment group $(1 = yes)$	(0.1609) -1.4051	(0.2915) $-2.1835$	(1.5510) -3.8596	(2.6338) -4.1760
Treatment group (1 — yes)	(2.8284)	(1.5515)	(3.2121)	(2.4014)
Post-9/11*Treatment group	1.6502	1.5125	2.5887	2.5460
, 0 1	(3.6481)	(2.8448)	(4.0488)	(3.4754)
Post-Iraq-War*Treatment group	0.1100	0.5257	0.2926	0.9300
	(3.8514)	(3.8048)	(4.2066)	(4.3146)
Post-Madrid-Bombings*Treatment group	0.5451	0.7195	2.5655	4.1732
D	(4.4748)	(4.4980)	(4.8605)	(4.6542)
Post-London-Bombings*Treatment group	-2.4442 (3.8860)	-1.5096 (3.7764)	-3.1504 (4.2015)	-3.7176 (3.8788)
Age in years	(3.8800)	0.7244***	(4.2013)	0.3376
rige in years		(0.0368)		(0.4634)
Age (squared)		-0.0088***		-0.0044
8 (* 1 * * * * *)		(0.0004)		(0.0061)
Married/cohabiting (1 = yes)		1.0451***		-0.6323
		(0.1396)		(1.1962)
No. of children under 16		-0.0867		-0.4795
- · · · · · · · · · · · · · · · · · · ·		(0.0630)		(0.5696)
Degree or equivalent $(1 = yes)$		-0.6059*** (0.1664)		1.8745
Higher education ( $1 = \text{yes}$ )		-0.4344*		(2.1407) $2.6676$
inglier education ( 1 — yes)		(0.1972)		(2.5306)
GCSE grades A*-C or equivalent ( $1 = yes$ )		0.0299		2.3335
,		(0.1539)		(3.4399)
Other qualification ( $1 = yes$ )		1.2549***		2.4253
		(0.1937)		(1.8334)
No qualification $(1 = yes)$		0.3473		3.6782
Tenure 3 to 6 months $(1 = yes)$		(0.2428)		(2.6051)
Tenure 3 to 6 months (1 = yes)		-0.1352 $(0.3973)$		0.5567 $(2.5175)$
Tenure 6 to 12 months $(1 = yes)$		-0.5943		-0.1974
renare o to 12 months (1 = yes)		(0.3595)		(2.3214)
Tenure 1 to 2 years $(1 = yes)$		-0.5517		0.9215
		(0.3333)		(2.1426)
Tenure 2 to 5 years $(1 = yes)$		-0.2457		0.6465
		(0.3116)		(2.1818)
Tenure 5 to 10 years $(1 = yes)$		-0.3003		3.4287
TF 10 + 00 (1 )		(0.3178)		(2.4899)
Tenure 10 to 20 years $(1 = yes)$		-0.7382*		3.3875
Tenure more then 20 years $(1 = yes)$		(0.3169) -0.9601**		(2.9076) $0.0344$
rendre more then 20 years (1 – yes)		(0.3316)		(3.5676)
Health problem hindering at work $(1 = yes)$		-0.7757***		-1.9734
		(0.1919)		(2.2088)
Firmsize $< 25$ employees $(1 = yes)$		0.4975***		0.7477
		(0.1281)		(1.3867)
Firmsize $> 500$ employees $(1 = yes)$		-0.5547***		0.1561
		(0.1519)		(1.3619)
Constant	43.0718***	24.7248***	45.5263***	46.8674***
I late Carl Contact	(0.1535)	(1.2406)	(1.4999)	(10.8510)
Industry fixed effects	no	yes	no	yes
Occupation fixed effects Region fixed effects	no no	yes	no no	yes
Region fixed effects Time fixed effects (months, years)	no no	yes yes	no no	yes yes
N	64.637	59,173	873	735

Table 20: Probability of employment, groups defined by current nationality, Probit-regression

TIONALITI, I RO				
	Treatment group: Current nationality from Arab / Muslim countr Comparison: British Comparison: Non-European			
	Compariso	on: British		omparison: Non-European
Variable	naive	adjusted	naive	adjusted
Observed after 9/11	-0.0443***	-0.0494*	-0.3159**	-0.1455
	(0.0120)	(0.0223)	(0.1033)	(0.1799)
Observed after 3/20/2003	-0.0239*	0.0417	0.1426	-0.1886
	(0.0102)	(0.0259)	(0.0833)	(0.2028)
Observed after 3/11/2004	-0.0135	0.0161	-0.0138	-0.0212
	(0.0109)	(0.0293)	(0.0872)	(0.2188)
Observed after 7/7/2005	-0.0558***	-0.0076	0.1008	-0.1296
	(0.0102)	(0.0225)	(0.0793)	(0.1642)
Treatment group $(1 = yes)$	-1.3067***	-1.5710***	-1.0408***	-0.8924***
, , ,	(0.1769)	(0.2023)	(0.1980)	(0.2298)
Post-9/11*Treatment group	0.2142	0.3903	0.4858*	0.4536
3 - 4	(0.2011)	(0.2346)	(0.2258)	(0.2629)
Post-Iraq-War*Treatment group	0.1184	0.0644	-0.0481	0.1087
	(0.1465)	(0.1826)	(0.1682)	(0.2023)
Post-Madrid-Bombings*Treatment group	-0.2781	-0.3623	-0.2778	-0.3899
	(0.1662)	(0.2058)	(0.1874)	(0.2323)
Post-London-Bombings*Treatment group	0.4386**	0.5285**	0.2820	0.3184
	(0.1600)	(0.1984)	(0.1783)	(0.2235)
Age in years	(0.1000)	0.1860***	(0.11.00)	0.0328
11go III yourb		(0.0023)		(0.0191)
Age (squared)		-0.0026***		-0.0005
Age (squared)		(0.0000)		(0.0003)
Married/cohabiting (1 = yes)		0.5758***		0.4697***
Married/contabiting (1 = yes)		(0.0108)		(0.0731)
No. of children under 16		-0.0900***		-0.0888**
No. of children under 10				
Degree or equivalent $(1 = yes)$		(0.0049) $0.0042$		(0.0315) -0.0050
Degree or equivalent ( 1 = yes)				
TT: 1 1 1: (1		(0.0135)		(0.1111)
Higher education ( $1 = yes$ )		0.0043		0.0200
GGGE 1 A*G : 1 · / 1		(0.0158)		(0.1466)
GCSE grades A*-C or equivalent ( $1 = yes$ )		-0.0704***		0.0706
0.1		(0.0120)		(0.1633)
Other qualification ( $1 = yes$ )		-0.1549***		-0.1283
		(0.0143)		(0.0968)
No qualification $(1 = yes)$		-0.4935***		-0.4495***
		(0.0156)		(0.1248)
Health problem hindering at work $(1 = yes)$		-0.9904***		-0.8073***
		(0.0109)		(0.1067)
Constant	-0.0284**	-2.8861***	-0.2944***	-1.7580**
	(0.0103)	(0.0971)	(0.0894)	(0.6736)
Occupation fixed effects	no	yes	no	yes
Region fixed effects	no	yes	no	yes
Time fixed effects (months, years)	no	yes	no	yes
N	141,219	126,278	3,360	2,702

Table 21: Weekly real wages, groups defined by religion, OLS-regression

	Comparison: Christians		Treatment group: Muslims Comparison: other Religion (not Sikh)	
Variable	naive	adjusted	naive	adjusted
Observed after 3/20/2003	-0.0205*	0.0133	0.1573	0.1407
	(0.0090)	(0.0141)	(0.0823)	(0.0992)
Observed after 3/11/2004	0.0150	0.0258	-0.0808	-0.0004
01 1 . 6 7 /7 /0005	(0.0085)	(0.0144)	(0.0900)	(0.1282)
Observed after $7/7/2005$	0.0106	0.0055	0.0367	0.0849
muslim	(0.0086) -0.1575*	(0.0112) -0.1209*	(0.0805) -0.1099	(0.0899) -0.0061
musmin	(0.0639)	(0.0493)	(0.0815)	(0.0698)
Post-Iraq-War*Treatment group	-0.0223	-0.0421	-0.2000	-0.1596
3	(0.0815)	(0.0633)	(0.1157)	(0.0920)
Post-Madrid-Bombings*Treatment group	-0.0401	-0.0381	0.0557	0.0432
	(0.0715)	(0.0531)	(0.1148)	(0.0847)
Post-London-Bombings*Treatment group	0.0808	0.1288*	0.0546	0.0285
	(0.0741)	(0.0513)	(0.1093)	(0.0805)
Age in years		0.0789***		0.0660***
A ( 1)		(0.0017)		(0.0116)
Age (squared)		-0.0009***		-0.0007***
Married/cohabiting $(1 = yes)$		(0.0000) 0.0782***		(0.0001) 0.0792
Married/conabiting (1 — yes)		(0.0058)		(0.0404)
No. of children under 16		0.0109***		-0.0271
1.01 of children ander 10		(0.0028)		(0.0160)
Degree or equivalent ( $1 = yes$ )		0.2220***		0.1626**
		(0.0076)		(0.0567)
Higher education ( $1 = yes$ )		0.0705***		0.0586
		(0.0083)		(0.0617)
GCSE grades A*-C or equivalent ( $1 = yes$ )		-0.0430***		-0.0772
0.1 10 11 (1		(0.0068)		(0.0616)
Other qualification ( $1 = yes$ )		-0.0590***		-0.0453
No qualification $(1 = yes)$		(0.0082) -0.1411***		(0.0594) -0.1490*
No qualification (1 — yes)		(0.0103)		(0.0626)
Tenure 3 to 6 months $(1 = yes)$		-0.0032		-0.0989
(- 'j')		(0.0215)		(0.1256)
Tenure 6 to 12 months $(1 = yes)$		-0.0060		-0.0033
,		(0.0189)		(0.0944)
Tenure 1 to 2 years $(1 = yes)$		0.0251		-0.0101
		(0.0182)		(0.0921)
Tenure 2 to 5 years $(1 = yes)$		0.0728***		0.0744
T (1)		(0.0170)		(0.0845)
Tenure 5 to 10 years $(1 = yes)$		0.0932***		0.0443
Tenure 10 to 20 years $(1 = yes)$		(0.0172) $0.1377***$		(0.0866)
Tenure 10 to 20 years $(1 = yes)$		(0.0171)		0.1672 (0.0890)
Tenure more then 20 years $(1 = yes)$		0.1687***		0.1674
Tendre more then 20 years (1 – yes)		(0.0176)		(0.0973)
Health problem hindering at work $(1 = yes)$		-0.0676***		-0.0868
3 - 1		(0.0084)		(0.0603)
Firmsize $< 25$ employees $(1 = yes)$		-0.0952***		-0.1667***
		(0.0059)		(0.0446)
Firmsize $> 500$ employees (1 = yes)		0.0824***		0.0937**
		(0.0057)	dodot	(0.0363)
Constant	6.0308***	3.9254***	5.9832***	4.4482***
	(0.0063)	(0.0489)	(0.0506)	(0.2963)
Industry fixed effects	no	yes	no	yes
Occupation fixed effects	no	yes	no	yes
Region fixed effects	no	yes	no	yes
Time fixed effects (months, years)  N	35,992	yes 32,907	no 1,032	909

Table 22: Hourly real wages, groups defined by religion, OLS-regression

	Comparison	: Christians	Treatment group: Muslims Comparison: other	s Religion (not Sikh)
Variable	naive	adjusted	naive	adjusted
Observed after 3/20/2003	-0.0216*	0.0075	0.1540	0.1198
	(0.0088)	(0.0130)	(0.0823)	(0.0956)
Observed after 3/11/2004	0.0209*	0.0186	-0.0380	-0.0388
Ob 1 -ft 7 /7 /2005	(0.0082)	(0.0136)	(0.0850)	(0.1212)
Observed after $7/7/2005$	0.0155 $(0.0082)$	0.0018 (0.0105)	-0.0075 (0.0752)	-0.0155 (0.0780)
muslim	-0.1564*	-0.1209**	-0.1455	-0.0131
	(0.0621)	(0.0469)	(0.0810)	(0.0659)
Post-Iraq-War*Treatment group	-0.0272	-0.0442	-0.2028	-0.1664
	(0.0773)	(0.0577)	(0.1127)	(0.0856)
Post-Madrid-Bombings*Treatment group	-0.0006	-0.0157	0.0582	0.0368
	(0.0633)	(0.0480)	(0.1058)	(0.0784)
Post-London-Bombings*Treatment group	0.0879	0.1404**	0.1109	0.0968
A	(0.0650)	(0.0488)	(0.0992)	(0.0731)
Age in years		0.0565***		0.0387***
Age (squared)		(0.0014) -0.0006***		(0.0102) -0.0004**
Age (squared)		(0.0000)		(0.0004
Married/cohabiting (1 = yes)		0.0550***		0.0732*
married, conditing (1 = yes)		(0.0054)		(0.0362)
No. of children under 16		0.0209***		-0.0143
		(0.0026)		(0.0140)
Degree or equivalent ( $1 = yes$ )		0.2590***		0.1510**
		(0.0074)		(0.0509)
Higher education ( $1 = yes$ )		0.0815***		0.0271
GGGT 1 1 4 G 1 1 1 1 1 1 1		(0.0081)		(0.0600)
GCSE grades A*-C or equivalent ( $1 = yes$ )		-0.0476***		-0.0655
Other qualification ( $1 = yes$ )		(0.0062) -0.0983***		(0.0554) -0.1030
Other quantication (1 = yes)		(0.0075)		(0.0526)
No qualification $(1 = yes)$		-0.1721***		-0.2254***
4 (-		(0.0093)		(0.0561)
Tenure 3 to 6 months $(1 = yes)$		0.0015		-0.0401
		(0.0182)		(0.1185)
Tenure 6 to 12 months $(1 = yes)$		-0.0005		0.0193
		(0.0160)		(0.0948)
Tenure 1 to 2 years $(1 = yes)$		0.0282		0.0750
m		(0.0155)		(0.0880)
Tenure 2 to 5 years $(1 = yes)$		0.0622***		0.0869
Tenure 5 to 10 years $(1 = yes)$		(0.0146) $0.0764***$		(0.0843) $0.0802$
Tenure 5 to 10 years $(1 = yes)$		(0.0149)		(0.0857)
Tenure 10 to 20 years $(1 = yes)$		0.1288***		0.2017*
10Hare 10 to 20 years (1 = yes)		(0.0148)		(0.0882)
Tenure more then 20 years $(1 = yes)$		0.1649***		0.1989*
3		(0.0153)		(0.0955)
Health problem hindering at work $(1 = yes)$		-0.0562***		-0.0356
		(0.0077)		(0.0551)
Firmsize $< 25$ employees $(1 = yes)$		-0.0931***		-0.1197**
T:		(0.0053)		(0.0395)
Firmsize $> 500$ employees (1 = yes)		0.0949***		0.1282***
Constant	2.3110***	(0.0055) $0.6567***$	2.3001***	(0.0343) 1.2248***
Constant	(0.0063)	(0.0399)	(0.0522)	(0.2934)
Industry fixed effects	(0.0003)	(0.0399) yes	no	(0.2934) yes
Occupation fixed effects	no	yes	no	yes
Region fixed effects	no	yes	no	yes
Time fixed effects (months, years)	no	yes	no	yes
N	35,992	32,907	1,032	909

Table 23: Weekly hours worked, groups defined by religion, OLS-regression

	Comparison		reatment group: Muslims Comparison: other	Religion (not Sikh)
Variable	naive	adjusted	naive	adjusted
Observed after 3/20/2003	0.4124*	0.4338	2.7409	3.4505
	(0.2096)	(0.4328)	(1.7329)	(3.6822
Observed after 3/11/2004	0.1716	0.2364	-3.1550	0.907
	(0.1884)	(0.4231)	(1.9564)	(2.9063
Observed after 7/7/2005	0.0649	0.0594	0.8361	3.8568
	(0.1800)	(0.3376)	(1.7862)	(2.7861
muslim	0.2077	0.7884	1.0096	0.909
	(1.2363)	(1.2195)	(1.6208)	(1.8313
Post-Iraq-War*Treatment group	-0.9413	-1.5409	-3.2698	-3.480
	(1.6740)	(1.6456)	(2.4044)	(2.6642)
Post-Madrid-Bombings*Treatment group	-0.7742	-0.1649	2.5524	2.381
	(1.5094)	(1.4553)	(2.4671)	(2.6076
Post-London-Bombings*Treatment group	0.8953	1.1415	0.1241	-0.260
	(1.5552)	(1.4892)	(2.3651)	(2.3857
Age in years		0.7524***		0.8690**
		(0.0479)		(0.3088
Age (squared)		-0.0091***		-0.0111**
		(0.0006)		(0.0038
Married/cohabiting (1 = yes)		0.8780***		0.737
		(0.1781)		(1.2245)
No. of children under 16		-0.1118		-0.696
		(0.0828)		(0.4103
Degree or equivalent ( $1 = yes$ )		-0.3710		-0.5999
		(0.2158)		(1.6624
Higher education ( $1 = yes$ )		-0.1988		-0.118
acan 1 14 a 1 1 (4 )		(0.2561)		(1.8302)
GCSE grades A*-C or equivalent ( $1 = yes$ )		0.1461		-0.7906
0.1		(0.1985)		(1.6990)
Other qualification ( $1 = yes$ )		1.1022***		1.5632
N. 110 (1 /1 )		(0.2453)		(1.7083)
No qualification $(1 = yes)$		0.2211		1.7928
Tenure 3 to 6 months $(1 = ves)$		(0.3149)		(1.9109)
Tenure 3 to 6 months $(1 = yes)$		0.2341		-5.8944
Tenure 6 to 12 months $(1 = yes)$		(0.5385)		(2.4922) 0.4122
Tenure 6 to 12 months (1 = yes)		0.2348		
Tenure 1 to 2 years $(1 = yes)$		(0.4827)		(2.1681)
Tenure 1 to 2 years (1 — yes)		0.0031		-3.8849
Tonura 2 to 5 years (1 - yes)		(0.4475)		(2.1850)
Tenure 2 to 5 years $(1 = yes)$		0.0132		-1.3720
Tenure 5 to 10 years $(1 = yes)$		(0.4180)		(1.9335) -1.4108
Tenure 5 to 10 years (1 — yes)		0.0829		(2.0330)
Tenure 10 to 20 years $(1 = yes)$		(0.4260) -0.3159		-3.6985
Tenure 10 to 20 years (1 = yes)		(0.4246)		(2.2117)
Tenure more then 20 years $(1 = yes)$		-0.5273		0.5535
Tenure more then 20 years (1 — yes)		(0.4385)		(2.4413)
Health problem hindering at work $(1 = yes)$		-0.7399**		-0.8862
reason problem influenting at work (1 — yes)		(0.2406)		(1.4463
Firmsize $< 25$ employees $(1 = yes)$				,
i ilmoize \ 20 employees (1 — yes)		0.2620 $(0.1672)$		-0.164 (1.2420
Firmsize $> 500$ employees $(1 = yes)$		-0.7335***		0.859
i illibize > 500 empioyees (1 — yes)		(0.1764)		(1.0910
Constant	42.3764***	20.8364***	41.5745***	32.6880**
Constant	(0.1530)	(1.4519)	(1.0541)	(12.6063
Industry fixed effects	(0.1550) no	(1.4519) yes	no	(12.0003 ye
Occupation fixed effects	no	-	no	
Region fixed effects	no	yes	no	ye
Time fixed effects (months, years)	no	yes yes	no	ye: ye:
N	37,844	34,578	1.092	957

Table 24: Probability of employment, groups defined by religion, Probit-regression

	Treatment group: Muslims				
	Comparison	: Christians	Comparison:	other Religion (not Sikh)	
Variable	naive	adjusted	naive	adjusted	
Observed after 3/20/2003	0.0202	0.0486	0.0587	0.3821*	
	(0.0137)	(0.0337)	(0.1071)	(0.1933)	
Observed after 3/11/2004	-0.0196	0.0026	-0.0183	0.2831	
	(0.0123)	(0.0344)	(0.1042)	(0.2062)	
Observed after 7/7/2005	-0.0625***	-0.0075	-0.0020	-0.2000	
	(0.0115)	(0.0262)	(0.0965)	(0.1572)	
muslim	-0.8410***	-0.9412***	-0.5717***	-0.3937**	
	(0.0589)	(0.0725)	(0.0951)	(0.1204)	
Post-Iraq-War*Treatment group	0.0116	0.0563	-0.0269	0.1624	
	(0.0801)	(0.0975)	(0.1330)	(0.1610)	
Post-Madrid-Bombings*Treatment group	0.0114	-0.0539	0.0101	-0.2479	
0 .	(0.0720)	(0.0883)	(0.1261)	(0.1531)	
Post-London-Bombings*Treatment group	0.0580	0.0623	-0.0025	0.0829	
	(0.0659)	(0.0817)	(0.1163)	(0.1421)	
Age in years	, ,	0.1894***	,	0.1095***	
		(0.0030)		(0.0164)	
Age (squared)		-0.0026***		-0.0015***	
3 ( 1 1 )		(0.0000)		(0.0002)	
Married/cohabiting (1 = ves)		0.5313***		0.4160***	
		(0.0140)		(0.0664)	
No. of children under 16		-0.0702***		-0.0780**	
to, or emidren ander 10		(0.0066)		(0.0238)	
Degree or equivalent ( $1 = yes$ )		0.0275		0.2392**	
Degree of equivalent (1 = yes)		(0.0177)		(0.0843)	
		-0.0245		0.0282	
		(0.0201)		(0.1086)	
GCSE grades A*-C or equivalent ( $1 = yes$ )		-0.1012***		-0.0865	
GODE grades it -e or equivalent ( 1 = yes)		(0.0158)		(0.0920)	
Other qualification ( $1 = ves$ )		-0.1844***		-0.3702***	
Other qualification (1 = yes)		(0.0182)		(0.0839)	
No qualification $(1 = yes)$		-0.5057***		-0.7108***	
(1 = yes)		(0.0205)		(0.0975)	
Health problem hindering at work $(1 = ves)$		-0.9588***		-0.9317***	
nearth problem initidering at work (1 = yes)		(0.0139)		(0.0713)	
Constant	-0.0928***	-2.4620***	-0.3621***	(0.0713) -1.7931***	
Constant	(0.0100)	(0.0973)	(0.0754)	(0.4465)	
Occupation fixed effects	(0.0100) no	(0.0973) yes	no	(0.4403) yes	
Region fixed effects	no	ves	no	yes	
Time fixed effects (months, years)	no	ves	no	yes	
N	84,733	75,686	4,700	3,972	

## **Working Paper Series in Economics**

(see www.leuphana.de/vwl/papers for a complete list)

No.69:	Maik Heinemann: E-stability and stability learning in models with asymmetric information, December 2007
No.68:	Joachim Wagner: Exporte und Produktivität in Industriebetrieben – Niedersachsen im
	interregionalen und internationalen Vergleich, Dezember 2007
No.67:	Stefan Baumgärtner and Martin F. Quaas: Ecological-economic viability as a criterion of strong sustainability under uncertainty, November 2007
No.66:	Kathrin Michael: Überbrückungsgeld und Existenzgründungszuschuss – Ergebnisse einer schriftlichen Befragung drei Jahre nach Gründungsbeginn, November 2007
No.65:	The International Study Group on Export and Productivity: Exports and Productivity – Comparable Evidence for 14 Countries, November 2007
No.64:	Lena Koller, Claus Schnabel und Joachim Wagner: Freistellung von Betriebsräten – Eine Beschäftigungsbremse?, November 2007
No.63:	Anne-Kathrin Last: The Monetary Value of Cultural Goods: A Contingent Valuation Study of the Municipal Supply of Cultural Goods in Lueneburg, Germany, October 2007
No.62:	Thomas Wein und Heike Wetzel: The Difficulty to Behave as a (regulated) Natural Monopolist – The Dynamics of Electricity Network Access Charges in Germany 2002 to 2005, September 2007
No.61:	Stefan Baumgärtner und Martin F. Quaas: Agro-biodiversity as natural insurance and the development of financial insurance markets, September 2007
No.60:	Stefan Bender, Joachim Wagner, Markus Zwick: KombiFiD - Kombinierte Firmendaten für Deutschland, September 2007
No.59:	Jan Kranich: Too much R&D? - Vertical differentiation in a model of monopolistic competition, August 2007
No.58:	Christian Papilloud und Ingrid Ott: Convergence or mediation? Experts of vulnerability and the vulnerability of experts' discourses on nanotechnologies – a case study, July 2007
No.57:	Ingrid Ott und Susanne Soretz: Governmental activity, integration and agglomeration, July 2007
No.56:	Nils Braakmann: Struktur und Erfolg von Ich-AG-Gründungen: Ergebnisse einer Umfrage im Arbeitsagenturbezirk Lüneburg, Juli 2007
No.55:	Nils Braakmann: Differences in the earnings distribution of self- and dependent employed German men – evidence from a quantile regression decomposition analysis, July 2007
No.54:	Joachim Waagner: Export entry, export exit, and productivity in German Manufacturing Industries, June 2007
	[forthcoming in: International Journal of the Economics of Business]
No.53:	Nils Braakmann: Wirkungen der Beschäftigungspflicht schwerbehinderter Arbeitnehmer
	<ul> <li>Erkenntnisse aus der Einführung des "Gesetzes zur Bekämpfung der Arbeitslosigkeit Schwerbehinderter", Juni 2007</li> </ul>

Jan Kranich und Ingrid Ott: Regionale Spitzentechnologie auf internationalen Märkten,

No.52:

Juni 2007

- No.51: Joachim Wagner: Die Forschungspotenziale der Betriebspaneldaten des Monatsberichts im Verarbeitenden Gewerbe, Mai 2007
   No.50: Stefan Baumgärtner, Frank Jöst und Ralph Winkler: Optimal dynamic scale and structure of a multi-pollution economy, May 2007
   No.49: Helmut Fryges und Joachim Wagner: Exports and productivity growth First evidence from a continuous treatment approach, May 2007
- No.48: *Ulrich Kaiser und Joachim Wagner:* Neue Möglichkeiten zur Nutzung vertraulicher amtlicher Personen- und Firmendaten, April 2007 [erscheint in: Perspektiven der Wirtschaftspolitik]
- No.47: Joachim Wagner: Jobmotor Mittelstand? Arbeitsplatzdynamik und Betriebsgröße in der westdeutschen Industrie, April 2007 [publiziert in: Vierteljahrshefte zur Wirtschaftsforschung, 76 (2007), 3, 76-87]
- No.46: *Christiane Clemens und Maik Heinemann:* Credit Constraints, Idiosyncratic Risks, and the Wealth Distribution in a Heterogenous Agent Model, March 2007
- No.45: *Jan Kranich:* Biotechnologie und Internationalisierung. Ergebnisse der Online-Befragung, März 2007
- No.44: *Joachim Wagner:* Entry, exit and productivity. Empirical results for German manufacturing industries, March 2007
- No.43: Joachim Wagner: Productivity and Size of the Export Market Evidence for West and East German Plants, 2004, March 2007
  [erscheint in: Jahrbücher für Nationalökonomie und Statistik]
- No.42: Joachim Wagner: Why more West than East German firms export, March 2007
- No.41: *Joachim Wagner:* Exports and Productivity in Germany, March 2007 [forthcoming in: Applied Economics Quarterly]
- No.40: Lena Koller, Klaus Schnabel und Joachim Wagner: Schwellenwerte im Arbeitsrecht.

  Höhere Transparenz und Effizienz durch Vereinheitlichung, Februar 2007

  [publiziert in: Perspektiven der Wirtschaftspolitik, 8 (2007), 3, 242-255]
- No.39: *Thomas Wein und Wiebke B. Röber:* Sind ausbildende Handwerksbetriebe erfolgreicher?, Januar 2007
- No.38: Joachim Wagner: Institut für Volkswirtschaft: Forschungsbericht 2006, Januar 2007
- No.37: *Nils Braakmann:* The impact of September 11<sup>th</sup>, 2001 on the job prospects of foreigners with Arab background Evidence from German labor market data, January 2007
- No.36: *Jens Korunig:* Regulierung des Netzmonopolisten durch Peak-load Pricing?, Dezember 2006
- No.35: *Nils Braakmann:* Die Einführung der fachkundigen Stellungnahme bei der Ich-AG, November 2006
- No.34: Martin F. Quaas and Stefan Baumgärtner: Natural vs. financial insurance in the management of public-good ecosystems, October 2006
  [forthcoming in: Ecological Economics]
- No.33: Stefan Baumgärtner and Martin F. Quaas: The Private and Public Insurance Value of Conservative Biodiversity Management, October 2006
- No.32: *Ingrid Ott and Christian Papilloud:* Converging institutions. Shaping the relationships between nanotechnologies, economy and society, October 2006 [published in: Bulletin of Science, Technology & Society 2007 (27), 4, 455-466]

- No.31: Claus Schnabel and Joachim Wagner: The persistent decline in unionization in western and eastern Germany, 1980-2004: What can we learn from a decomposition analysis?, October 2006
  [published in: Industrielle Beziehungen/The German Journal of Industrial Relations 14 (2007), 118-132]
- No.30: Ingrid Ott and Susanne Soretz: Regional growth strategies: fiscal versus institutional governmental policies, September 2006
  [forthcoming in: Economic Modelling]
- No.29: Christian Growitsch and Heike Wetzel: Economies of Scope in European Railways: An Efficiency Analysis, July 2006
- No.28: Thorsten Schank, Claus Schnabel and Joachim Wagner: Do exporters really pay higher wages? First evidence from German linked employer-employee data, June 2006 [published in in: Journal of International Economics 72 (2007), 1, 52-74]
- No.27: Joachim Wagner: Markteintritte, Marktaustritte und Produktivität Empirische Befunde zur Dynamik in der Industrie, März 2006 [erscheint in: Allgemeines Statistisches Archiv, Heft 3/2007]
- No.26: Ingrid Ott and Susanne Soretz: Governmental activity and private capital adjustment,

  March 2006

  [forthcoming in: Icfai Journal of Managerial Economics]
- No.25: Joachim Wagner: International Firm Activities and Innovation:

  Evidence from Knowledge Production Functions for German Firms, March 2006
- No.24: Ingrid Ott und Susanne Soretz: Nachhaltige Entwicklung durch endogene
  Umweltwahrnehmung, März 2006
  publiziert in: Clemens, C., Heinemann, M. & Soretz, S., Auf allen Märkten zu Hause
  (Gedenkschrift für Franz Haslinger), Marburg: Metropolis, 2006, 233-256
- No.23: John T. Addison, Claus Schnabel, and Joachim Wagner: The (Parlous) State of German Unions, February 2006
  [forthcoming in: Journal of Labor Research 28 (2007), 3-18]
- No.22: Joachim Wagner, Thorsten Schank, Claus Schnabel, and John T. Addison: Works Councils, Labor Productivity and Plant Heterogeneity: First Evidence from Quantile Regressions, February 2006

  [published in: Jahrbücher für Nationalökonomie und Statistik 226 (2006), 505 518]
- No.21: Corinna Bunk: Betriebliche Mitbestimmung vier Jahre nach der Reform des BetrVG: Ergebnisse der 2. Befragung der Mitglieder des Arbeitgeberverbandes Lüneburg Nordostniedersachsen, Februar 2006
- No.20: Jan Kranich: The Strength of Vertical Linkages, July 2006
- No.19: Jan Kranich und Ingrid Ott: Geographische Restrukturierung internationaler
  Wertschöpfungsketten Standortentscheidungen von KMU aus regionalökonomischer
  Perspektive, Februar 2006
- No.18: Thomas Wein und Wiebke B. Röber: Handwerksreform 2004 Rückwirkungen auf das Ausbildungsverhalten Lüneburger Handwerksbetriebe?, Februar 2006
- No.17: Wiebke B. Röber und Thomas Wein: Mehr Wettbewerb im Handwerk durch die Handwerksreform?. Februar 2006

- No.16: Joachim Wagner: Politikrelevante Folgerungen aus Analysen mit wirtschaftsstatistischen Einzeldaten der Amtlichen Statistik, Februar 2006 [publiziert in: Schmollers Jahrbuch 126 (2006) 359-374]
- No.15: Joachim Wagner: Firmenalter und Firmenperformance
  Empirische Befunde zu Unterschieden zwischen jungen und alten Firmen
  in Deutschland, September 2005
  [publiziert in: Lutz Bellmann und Joachim Wagner (Hrsg.), Betriebsdemographie
  (Beiträge zur Arbeitsmarkt- und Berufsforschung, Band 305), Nürnberg: IAB der BA,
  83-111]
- No.14: Joachim Wagner: German Works Councils and Productivity:

  First Evidence from a Nonparametric Test, September 2005

  [forthcoming in: Applied Economics Letters]
- No.13: Lena Koller, Claus Schnabel und Joachim Wagner: Arbeitsrechtliche Schwellenwerte und betriebliche Arbeitsplatzdynamik: Eine empirische Untersuchung am Beispiel des Schwerbehindertengesetzes, August 2005

  [publiziert in: Zeitschrift für ArbeitsmarktForschung/ Journal for Labour Market Research 39 (2006), 181-199]
- No.12: Claus Schnabel and Joachim Wagner: Who are the workers who never joined a union? Empirical evidence from Germany, July 2005
  [published in: Industrielle Beziehungen/ The German Journal of Industrial Relations 13 (2006), 118-131]
- No.11: Joachim Wagner: Exporte und Produktivität in mittelständischen Betrieben Befunde aus der niedersächsischen Industrie (1995 2004), June 2005 [publiziert in: Niedersächsisches Landesamt für Statistik, Statistische Berichte Niedersachsen, Sonderausgabe: Tagung der NLS am 9. März 2006, Globalisierung und regionale Wirtschaftsentwicklung Datenlage und Datenbedarf in Niedersachsen. Hannover, Niedersächsisches Landesamt für Statistik, Juli 2006, 18 29]
- No.10: Joachim Wagner: Der Noth gehorchend, nicht dem eignen Trieb.

  Nascent Necessity and Opportunity Entrepreneurs in Germany.

  Evidence from the Regional Entrepreneurship Monitor (REM), May 2005

  [published in: RWI: Mitteilungen. Quarterly 54/ 55 (2003/04), 287-303

  {published June 2006}]
- No. 9: Gabriel Desgranges and Maik Heinemann: Strongly Rational Expectations Equilibria with Endogenous Acquisition of Information, March 2005
- No. 8: Joachim Wagner: Exports, Foreign Direct Investment, and Productivity: Evidence from German Firm Level Data, March 2005
  [published in: Applied Economics Letters 13 (2006), 347-349]
- No. 7: Thomas Wein: Associations' Agreement and the Interest of the Network Suppliers The Strategic Use of Structural Features, March 2005
- No. 6: Christiane Clemens and Maik Heinemann: On the Effects of Redistribution on Growth and Entrepreneurial Risk-Taking, March 2005
- No. 5: Christiane Clemens and Maik Heinemann: Endogenous Redistributive Cycles An overlapping Generations Approach to Social Conflict and Cyclical Growth, March 2005

- No. 4: Joachim Wagner: Exports and Productivity: A Survey of the Evidence from Firm Level
  Data, March 2005
  [published in: The World Economy 30 (2007), 1, 60-82]
- No. 3: Thomas Wein and Reimund Schwarze: Is the Market Classification of Risk Always Efficient? Evidence from German Third Party Motor Insurance, March 2005
- No. 2: Ingrid Ott and Stephen J. Turnovsky: Excludable and Non-Excludable Public Inputs: Consequences for Economic Growth, June 2005 (Revised version) [published in: Economica 73 (2006), 292, 725-742 also published as CESifo Working Paper 1423]
- No. 1: Joachim Wagner: Nascent and Infant Entrepreneurs in Germany.

  Evidence from the Regional Entrepreneurship Monitor (REM), March 2005

  [published in: Simon C. Parker (Ed.), The Life Cycle of Entrepreneurial Ventures

  (International Handbook Series on Entrepreneurship, Volume 3), New York etc.: Springer, 2006, 15-37]

Leuphana Universität Lüneburg Institut für Volkswirtschaftslehre Postfach 2440 D-21314 Lüneburg

Tel.: ++49 4131 677 2321 email: brodt@leuphana.de

www.leuphana.de/vwl/papers